

# Sub-logarithmic fluctuations for internal DLA

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## Abstract

We consider internal diffusion limited aggregation in dimension larger or equal to two. This is a random cluster growth model, where random walks start at the origin of the  $d$ -dimensional lattice, one at a time, and stop moving when reaching a site not occupied by previous walks. It is known that the asymptotic shape of the cluster is a sphere. When dimension is two or more, we have shown that the inner (resp. outer) fluctuations of its radius is at most of order  $\log(\text{radius})$  (resp.  $\log^2(\text{radius})$ ). Using the same approach, we improve the upper bound on the inner fluctuation to  $\sqrt{\log(\text{radius})}$  when  $d$  is larger than or equal to three. The inner fluctuation is then used to obtain a similar upper bound on the outer fluctuation.

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## 1 Introduction

This note is a companion to our paper [1]. There, we introduced a family of cluster growth models with a spherical asymptotic shape but a wide diversity of shape fluctuations. Internal diffusion limited aggregation (internal DLA) was one member of this family. More precisely, the internal DLA cluster of volume  $N$ , say  $A(N)$ , is obtained inductively as follows. Initially, we assume that the explored region is empty, that is  $A(0) = \emptyset$ . Then, consider  $N$  independent discrete-time random walks  $S_1, \dots, S_N$  starting from 0. Assume  $A(k-1)$  is obtained, and define

$$\tau_k = \inf \{t \geq 0 : S_k(t) \notin A(k-1)\}, \quad \text{and} \quad A(k) = A(k-1) \cup \{S_k(\tau_k)\}. \quad (1.1)$$

We call explorers the random walks obeying the aggregation rule (1.1). We say that the  $k$ -th explorer is *settled* on  $S_k(\tau_k)$  after time  $\tau_k$ , and is *unsettled* before time  $\tau_k$ . The cluster  $A(N)$  is interpreted as the positions of the  $N$  settled explorers.

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In this paper we show how the tools developed in [1] lead in dimension  $d \geq 3$  to sharper estimates on the fluctuations of  $A(N)$  with respect to its spherical asymptotic shape. We keep the notation of [1], and recall some to make the paper as self-contained as possible. We denote with  $\|\cdot\|$  the euclidean norm on  $\mathbb{R}^d$ . For any  $x$  in  $\mathbb{R}^d$  and  $r$  in  $\mathbb{R}$ , set

$$B(x, r) = \{y \in \mathbb{R}^d : \|y - x\| < r\} \quad \text{and} \quad \mathbb{B}(x, r) = B(x, r) \cap \mathbb{Z}^d. \quad (1.2)$$

For  $\Lambda \subset \mathbb{Z}^d$ ,  $|\Lambda|$  denotes the number of sites in  $\Lambda$ , and the boundary of  $\Lambda$  is  $\partial\Lambda = \{z \notin \Lambda : \exists y \in \Lambda, \|y - z\| = 1\}$ . For a simple random walk, let  $H(\Lambda)$  denotes its first hitting time of  $\Lambda$ . The inner error  $\delta_I(n)$  is such that

$$n - \delta_I(n) = \sup \{r \geq 0 : \mathbb{B}(0, r) \subset A(|\mathbb{B}(0, n)|)\}. \quad (1.3)$$

Also, the outer error  $\delta_O(n)$  is such that

$$n + \delta_O(n) = \inf \{r \geq 0 : A(|\mathbb{B}(0, n)|) \subset \mathbb{B}(0, r)\}. \quad (1.4)$$

Our main result is as follows.

**Proposition 1.1** *There are constants  $\{\alpha_d, \beta_d, d \geq 3\}$  such that in dimension  $d \geq 3$ , with probability 1,*

$$\limsup \frac{\delta_I(n)}{\sqrt{\log(n)}} \leq \alpha_d, \quad \text{and} \quad \limsup \frac{\delta_O(n)}{\sqrt{\log(n)}} \leq \beta_d. \quad (1.5)$$

**Remark 1.2** *For  $d = 2$ , we also show that there are constants  $\alpha_2, \beta_2$  such that, with probability 1,*

$$\limsup \frac{\delta_I(n)}{\log(n)} \leq \alpha_2, \quad \text{and} \quad \limsup \frac{\delta_O(n)}{\log(n)} \leq \beta_2. \quad (1.6)$$

*Such a bound on the inner error was already obtained in [1]. Recently, Jerison, Levine, and Sheffield [2] have established, in dimension two and with a different method, the same estimate on both the inner and outer errors.*

Let us describe the main steps. The inner error is at the heart of the argument. It is based on a large deviation estimate which refines our previous estimates, with interest of its own. For a real  $x$ , let  $\lfloor x \rfloor$  be the integer part of  $x$ .

**Lemma 1.3** *Choose  $R$  large enough, and  $A \geq 1$ . Assume that  $\lfloor AR^d \rfloor$  explorers lie initially on  $\mathbb{B}(0, R/2)$ . We call  $\eta$  the initial configuration of these explorers and  $A(\eta)$  the cluster they produce. There are positive constants  $\{\kappa_d, d \geq 2\}$  independent of  $R$  and  $A$ , such that when  $d \geq 3$*

$$P(\mathbb{B}(0, R) \not\subset A(\eta)) \leq \exp(-\kappa_d AR^2), \quad (1.7)$$

*and when  $d = 2$ , we have*

$$P(\mathbb{B}(0, R) \not\subset A(\eta)) \leq \exp\left(-\kappa_2 \frac{AR^2}{\log(R)}\right). \quad (1.8)$$

**Remark 1.4** *The reason behind the previous Lemma, in  $d \geq 3$ , is that out of  $\lfloor AR^d \rfloor$  explorers, only about  $AR^2$  eventually hit a fixed site on the boundary of  $\mathbb{B}(0, R)$ , so that it is only these very explorers that need to be pushed away from this very site if we want it unoccupied. The cost should be proportional to  $AR^2$ .*

For the outer error, we use a large deviation estimate symmetrical to Lemma 1.3 as well as our coupling between internal DLA and the *flashing process* of [1]. This last large deviation estimate was recently proved by Jerison, Levine and Sheffield in [2].

**Lemma 1.5** *[Lemma A of Jerison et al. [2]] For  $\beta$ , and  $R$  positive reals, assume that  $\lfloor \beta R^d \rfloor$  explorers lie initially outside  $\mathbb{B}(0, R)$ . We call  $\eta$  the initial configuration of these explorers and  $A(\eta)$  the cluster they produce. There are positive constants  $\{\kappa'_d, d \geq 2\}$ , such that for  $\beta$  small enough, we have when  $d \geq 3$ ,*

$$P(0 \in A(\eta)) \leq \exp(-\kappa'_d R^2), \tag{1.9}$$

whereas when  $d = 2$ , we have

$$P(0 \in A(\eta)) \leq \exp\left(-\kappa'_2 \frac{R^2}{\log(R)}\right). \tag{1.10}$$

We give an alternative proof of this result, based on estimating the probability of crossing a shell, while avoiding traps.

**Lemma 1.6** *Consider  $d \geq 2$ . Fix a positive real  $R$ , and start a random walk on  $z \in \partial\mathbb{B}(0, 2R)$ . There are positive constants  $\{\kappa_d, a_d\}$  such that for any  $V$  subset of the shell  $\mathcal{S} = \mathbb{B}(0, 2R) \setminus \mathbb{B}(0, R)$ , we have*

$$P_z(H(\mathbb{B}(0, R)) < H(V^c)) \leq \exp\left(a_d - \kappa_d \frac{R}{\rho}\right) \quad \text{where} \quad \rho^{d-1} = \frac{|V|}{R}. \tag{1.11}$$

**Remark 1.7**  $V^c = \mathcal{S} \setminus V$  is interpreted as traps. Note that  $\rho$  is proportional to the radius of a cylinder of height  $R$  and volume  $|V|$ . We can also read (1.11) in the following way.

$$P_z(H(\mathbb{B}(0, R)) < H(V^c)) \leq \exp\left(a_d - \kappa_d \left(\frac{R^d}{|V|}\right)^{\frac{1}{d-1}}\right). \tag{1.12}$$

This shows that for (1.12) to be an effective inequality, one needs that  $|V|$  be smaller than  $R^d$ . The power  $1/(d-1)$  on  $R^d/V$  in (1.12) is not important in proving (1.5). If one were willing to accept the weaker power  $1/d$ , then one would have the following simple heuristics in  $d \geq 3$ . Let  $t$  denotes the time the walk spends in the annulus of height  $R$ . On one hand, the central limit scaling yields that the probability of such a stay is of order  $\exp(-cR^2/t)$ . On the other hand, all this time should be spent on sites of  $V$ , and it is well known that this event has probability of order  $\exp(-\kappa_d t/|V|^{2/d})$ . Putting together these opposite requirements, and optimizing over  $t$ , we find a statement weaker than (1.12), but enough for our present purpose.

$$P_z(H(\mathbb{B}(0, R)) < H(V^c)) \leq \exp\left(a_d - \kappa_d \left(\frac{R^d}{|V|}\right)^{\frac{1}{d}}\right). \tag{1.13}$$

Even though inequality (1.13) is not written in [1], it was the motivation behind the introduction of flashing processes in [1], which were basically used to bypass this type of estimate.

The rest of the paper is organized as follows. In Section 2, we enounce some known results: we recall the approach of Lawler, Bramson and Griffeath [3], and useful large deviation estimates. Then, the inner error estimate is proved in Section 3. In Section 4, we show how a flashing process permits a simple control on the outer error. Finally, we have gathered in an Appendix the proof of the large deviations Lemmas 1.3, 1.5 and 1.6.

## 2 Prerequisites

### 2.1 Notation

We recall some notation of [1]. The state space of configurations is  $\mathbb{N}^{\mathbb{Z}^d}$ , and its elements are denoted  $\eta$ , and represents starting conditions for a set of explorers, or random walks. Two types of initial configurations play an important role here: (i) the configuration  $n\mathbf{1}_{z^*}$  formed by  $n$  trajectories starting on a given site  $z^*$ , (ii) for  $\Lambda \subset \mathbb{Z}^d$ , the configuration  $\mathbf{1}_\Lambda$  that we simply identify with  $\Lambda$ . For any configuration  $\eta \in \mathbb{N}^{\mathbb{Z}^d}$  we write

$$|\eta| = \sum_{z \in \mathbb{Z}^d} \eta(z). \quad (2.1)$$

**Definition 2.1** *Let  $R \in \mathbb{R}_+ \cup \{\infty\}$ . For  $z \in \mathbb{B}(0, R) \cup \partial\mathbb{B}(0, R)$ , we denote by  $M_R(\eta, z)$  (resp.  $W_R(\eta, z)$ ) the number of simple random walks (resp. explorers) initially on  $\eta$  that hit  $z$  when or before exiting  $\mathbb{B}(0, R)$ . Thus, when  $z \in \partial\mathbb{B}(0, R)$ ,  $M_R(\eta, z)$  (resp.  $W_R(\eta, z)$ ) is the number of simple random walks (resp. explorers) which exit  $\mathbb{B}(0, R)$  exactly on  $z$ .*

As in [5] (Section 3), it is useful to stop explorers as they reach  $\partial\mathbb{B}(0, R)$ , for some  $R > 0$ , then to define  $A_R(\eta)$  as the set of positions of settled explorers.

**Definition 2.2** *Consider  $R \in \mathbb{R} \cup \{\infty\}$ . We set*

$$\forall z \in \mathbb{B}(0, R), \quad \tilde{M}_R(\eta, z) = W_R(\eta, z) + M_R(A_R(\eta), z). \quad (2.2)$$

Finally, for any function  $F : \mathbb{Z}^d \rightarrow \mathbb{R}$ , and  $\Lambda \subset \mathbb{Z}^d$ , we denote

$$F(\Lambda) = \sum_{z \in \Lambda} F(z).$$

### 2.2 On a classical approach

We recall the approach of Lawler, Bramson and Griffeath in [3]. Send  $N = |\mathbb{B}(0, n)|$  explorers from the origin. The approach of [3] is based on the following observations. (i) If explorers

would not settle, they would just be independent random walks; (ii) exactly one explorer occupies each site of the cluster. Then, observations (i) and (ii) imply that for any  $z \in \mathbb{B}(0, n)$

$$\tilde{M}_n(N \mathbb{I}_0, z) := W_n(N \mathbb{I}_0, z) + M_n(A_n(N), z) \stackrel{\text{law}}{=} M_n(N \mathbb{I}_0, z). \quad (2.3)$$

Note that for any set  $\Lambda \subset \mathbb{B}(0, n)$ ,  $M_n(\Lambda, z)$  is a sum of independent Bernoulli variables. Note also that  $A_n(N) \subset \mathbb{B}(0, n)$  so that

$$W_n(N \mathbb{I}_0, z) + M_n(\mathbb{B}(0, n), z) \geq \tilde{M}_n(N \mathbb{I}_0, z). \quad (2.4)$$

However, Lawler et al. did not use that  $W_n(N \mathbb{I}_0, z)$  and  $M_n(\mathbb{B}(0, n), z)$  were independent. They could only obtain a rough estimate on the lower tail of  $W_n(N \mathbb{I}_0, z)$ . This in turn gave some estimates on the inner error, which was used to derive bounds on the outer error, by using that the cluster covers  $\mathbb{B}(0, n - \delta_I(n))$ . In other words, from (2.3), and the definition of  $\delta_I(n)$

$$W_\infty(N \mathbb{I}_0, z) + M_\infty(\mathbb{B}(0, n - \delta_I(n)), z) \leq \tilde{M}_\infty(N \mathbb{I}_0, z).$$

Therefore, if  $\delta_I(n)$  is likely to be smaller than  $R$ ,

$$\mathbb{1}_{\{\delta_I(n) < R\}} (W_\infty(N \mathbb{I}_0, z) + M_\infty(\mathbb{B}(0, n - \delta_I(n)), z)) \leq \tilde{M}_\infty(N \mathbb{I}_0, z). \quad (2.5)$$

Let us now recall a simple tool of [1] in estimating deviations in view of (2.4), and (2.5).

## 2.3 On sums of Bernoulli variables

We first enounce the lower tail estimate.

**Lemma 2.3** *Suppose that a sequence of random variables  $\{W_n, M_n, L_n, \tilde{M}_n \mid n \in \mathbb{N}\}$ , and a sequence of real numbers  $\{c_n, n \in \mathbb{N}\}$ , satisfy for each  $n \in \mathbb{N}$*

$$W_n + L_n + c_n \geq \tilde{M}_n, \quad \text{and} \quad \tilde{M}_n \stackrel{\text{law}}{=} M_n. \quad (2.6)$$

*Assume that  $W_n$  and  $L_n$  are independent, and that  $L_n$  and  $M_n$  both are sums of independent Bernoulli variables. Assume that the Bernoulli variables  $\{Y_1^{(n)}, \dots, Y_{N_n}^{(n)}\}$  summing up to  $L_n$ , satisfy for some  $\kappa > 1$*

$$(H1) \quad \sup_n \sup_{i \leq N_n} E[Y_i^{(n)}] < \frac{\kappa - 1}{\kappa},$$

$$(H2) \quad \mu_n := E[M_n] - E[L_n] \geq 0.$$

*Then, for any  $n$  in  $\mathbb{N}$  and  $\xi_n$  in  $\mathbb{R}$  we have for all  $\lambda \geq 0$*

$$P(W_n < \xi_n) \leq \exp \left( -\lambda(\mu_n - \xi_n - c_n) + \frac{\lambda^2}{2} \left( \mu_n + \kappa \sum_{i=1}^{N_n} E[Y_i^{(n)}]^2 \right) \right). \quad (2.7)$$

The upper tail estimate needs other assumptions.

**Lemma 2.4** *Assume for each  $n \in \mathbb{N}$ , and for an event  $\mathcal{A}_n$*

$$\mathbb{I}_{\mathcal{A}_n}(W_n + L_n) \leq \tilde{M}_n, \quad \text{and} \quad \tilde{M}_n \stackrel{\text{law}}{=} M_n. \quad (2.8)$$

*Assume that  $W_n$  and  $L_n$  are independent,  $\mathbb{I}_{\mathcal{A}_n}$  and  $L_n$  are independent, and that  $L_n$  and  $M_n$  both are sums of independent Bernoulli variables such that  $\mu_n := E[M_n] - E[L_n] \geq 0$ . Then, for all  $n$  in  $\mathbb{N}$ ,  $\xi_n$  in  $\mathbb{R}$  and  $\lambda \in [0, \log 2]$ ,*

$$P(W_n \geq \xi_n, \mathcal{A}_n) \leq \exp \left( -\lambda(\xi_n - \mu_n) + \lambda^2 \left( \mu_n + 4 \sum_i E[Y_i^{(n)}]^2 \right) \right). \quad (2.9)$$

**Remark 2.5** *This lower (resp. upper) tail estimate turns out to be useful when  $\xi_n + c_n$  is less than (resp.  $\xi_n$  is more than)  $E[M_n] - E[L_n]$ . By Lemmas 2.3 and 2.4 tail estimates reduce to a three step strategy: i) estimation of  $E[M_n] - E[L_n]$  ii) estimation of  $\sum_i E^2[Y_i^{(n)}]$  iii) optimization in  $\lambda$ . We emphasize that this strategy does not require any control of the variance of  $W_n$ .*

**Proof of Lemmas 2.3 and 2.4:** As in [1] this is an application of Lemma 2.2 of [1]. For the lower tail, using the exponential Chebishev inequality, the independence between  $W_n$  and  $L_n$ , formula (2.6) and after centering the random variables, we get

$$P(W_n < \xi_n) \leq \frac{E[e^{-\lambda(M_n - E[M_n])}]}{E[e^{-\lambda(L_n - E[L_n])}]} e^{-\lambda(E[M_n] - E[L_n] - \xi_n - c_n)}. \quad (2.10)$$

If we define, for  $t \in \mathbb{R}$ ,  $f(t) = e^t - (1 + t)$  and  $g(t) = (e^t - 1)^2$ , then Lemma 2.2 of [1] yields

$$\frac{E[e^{-\lambda(M_n - E[M_n])}]}{E[e^{-\lambda(L_n - E[L_n])}]} \leq \exp \left\{ f(-\lambda)(E[M_n] - E[L_n]) + \frac{\kappa}{2} g(-\lambda) \sum_{i=1}^{N_n} E^2[Y_i^{(n)}] \right\}. \quad (2.11)$$

We conclude by observing that for all  $t \in \mathbb{R}$ ,

$$f(t) \leq \frac{t^2}{2} e^{[t]_+} \quad \text{and} \quad g(t) \leq t^2 e^{2[t]_+}, \quad (2.12)$$

where  $[\cdot]_+$  stands for the positive part. The proof for the upper tail is similar.

## 2.4 On a discrete mean value property of Green's function

We recall now the key result behind the sphericity of the asymptotic shape.

**Proposition 2.6** *Consider  $d \geq 2$ . There is a constant  $K_d$  such that, for any  $n$  and  $R$  with  $n - n^{\frac{1}{3}} \geq R \geq n$  and  $z$  in  $\mathbb{B}(0, R)$  with  $n - \|z\| \geq 1$ ,*

$$\left| |\mathbb{B}(0, R)| G_n(0, z) - \sum_{y \in \mathbb{B}(0, R)} G_n(y, z) \right| \leq K_d. \quad (2.13)$$

**Proof :** For  $n - R$  large enough this is Proposition 4.2 of [1]. For  $n = R$  this is a direct consequence of Lemmas 2 and 3 of [5]. For the remaining cases one can use the same Lemmas in conjunction with Lemma 5 of [5].

**Remark 2.7** For the inner bound we will use Proposition 2.6 with  $R = n$ . For the outer bound we will use Proposition 2.6 with  $n - R$  of order  $\log n$  in dimension 2 and  $\sqrt{\log n}$  in dimension  $d \geq 3$ .

### 3 Inner error

#### 3.1 Exploration by waves

We choose the following height sequence. For any positive integer  $n$ ,  $h(n) = \sqrt{\log(n)}$  in  $d \geq 3$ , and  $h(n) = \log(n)$  in  $d = 2$ . We partition  $\mathbb{Z}^d$  into concentric shells of heights  $h(n)$ . We define  $\mathcal{S}_0 = \mathbb{B}(0, h(n))$ , and for  $k \geq 1$ ,

$$\mathcal{S}_k = \mathbb{B}(0, (k + 1)h(n)) \setminus \mathbb{B}(0, kh(n)), \quad \text{and} \quad \Sigma_k = \partial\mathbb{B}(0, kh(n)). \quad (3.1)$$

We realize the internal DLA as an exploration wave process, where concentric shells are covered in turn (see Section 3 of [5]).

We fix an integer  $k$ . For a site  $z \in \Sigma_k$ , we call *cell* centered on  $z$ ,  $\mathcal{C}(z) := \mathbb{B}(z, h(n)) \cap \mathcal{S}_k$ , and we call *tile* centered on  $z$ ,  $\mathcal{T}(z) := \mathbb{B}(z, h(n)/2) \cap \Sigma_k$ . A generic cell is denoted  $\mathcal{C}$ , and a generic tile is denoted  $\mathcal{T}$ . Note the obvious facts

$$\bigcup_{z \in \Sigma_k} \mathbb{B}(z, h(n)) \supset \mathcal{S}_k. \quad (3.2)$$

Before covering shell  $\mathcal{S}_k$ , one stops the unsettled explorers on  $\Sigma_k$ . Following [1], for  $z \in \Sigma_k$ , we prove that the  $W_{kh(n)}(N \mathbb{I}_0, \mathcal{T})$  explorers stopped on  $\mathcal{T} = \mathcal{T}(z)$  are likely to cover  $\mathcal{C}(z)$ , if  $kh(n) \leq n - Ah(n)$  for a large enough constant  $A$ . As first observed in [3]

$$W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) + M_{kh(n)}(\mathbb{B}(0, kh(n)), \mathcal{T}) \geq \tilde{M}_{kh(n)}(N \mathbb{I}_0, \mathcal{T}). \quad (3.3)$$

Since (3.3) corresponds to an inequality of type (2.6), we wish to use Lemma 2.3, but we need to ensure (H1) and (H2).

First, if  $\tilde{\mathbb{B}}(r)$  denotes the sites of  $\mathbb{B}(0, kh(n))$  at a distance less than  $r$  from  $\mathcal{T}$ , there is  $L$  and  $\rho_d > 1$ , (which depend only on dimension) such that

$$\sup_{y \in \mathbb{B}(0, kh(n)) \setminus \tilde{\mathbb{B}}(Lh(n))} P_y(S(H(\partial\Sigma_k)) \in \mathcal{T}) < \frac{\rho_d - 1}{\rho_d} \quad (3.4)$$

(see Lemma 4.5 of [1]). Set  $c_n = |\tilde{\mathbb{B}}(Lh(n))|$ , and note that  $c_n \leq c(Lh(n))^d$  for some constant  $c$ . From (3.3) we have

$$W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) + M_{kh(n)}(\mathbb{B}(0, kh(n)) \setminus \tilde{\mathbb{B}}(Lh(n)), \mathcal{T}) \geq \tilde{M}_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) - c_n. \quad (3.5)$$

We will use Lemma 2.3 with  $L_n = M_{kh(n)}(\mathbb{B}(0, kh(n)) \setminus \tilde{\mathbb{B}}(Lh(n)), \mathcal{T})$  and we note that (H1) is ensured by (3.4). Let us define

$$\mu = E [M_{kh(n)}(N \mathbb{I}_0, \mathcal{T})] - E [M_{kh(n)}(\mathbb{B}(0, kh(n)) \setminus \tilde{\mathbb{B}}_{\mathcal{T}}(Lh(n)), \mathcal{T})]. \quad (3.6)$$

We consider the event that  $\mathcal{S}_k$  is not covered, and use the bound

$$\begin{aligned} P(\mathcal{S}_k \text{ not covered}) &\leq P\left(\exists \mathcal{T} \subset \Sigma_k : W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) < \frac{1}{3}\mu\right) \\ &\quad + P\left(\mathcal{S}_k \text{ not covered}, \forall \mathcal{T} \subset \Sigma_k : W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) \geq \frac{1}{3}\mu\right). \end{aligned} \quad (3.7)$$

In the next Sections, we compute  $\mu$ , and estimate the probabilities of events on the right hand side of (3.7).

### 3.1.1 Mean number of explorers crossing a tile

When  $\mathcal{T}$  is the tile of a cell  $\mathcal{C}$ , which has side  $h(n)$ , and belongs to shell  $\mathcal{S}_k \subset \mathbb{B}(0, n)$ , at a distance  $Ah(n)$  from  $\mathbb{B}(0, n)$ , then we have that for some positive constants  $\{c_d, d \geq 2\}$

$$\mu \geq c_d Ah(n)^d. \quad (3.8)$$

Indeed, this follows as in [1], Section 5.2 and relies on Proposition 2.6. Note that (3.8) ensures (H2).

### 3.1.2 $W_{kh(n)}(N \mathbb{I}_0, \mathcal{T})$ is unlikely to be small

By the computations (5.25) and (5.26) of Section 5.2 of [1], there are constants  $C_d$  such that

$$\sum_{y \in \mathbb{B}(0, kh(n)) \setminus \tilde{\mathbb{B}}(Lh(n))} P_y^2(S(H(\partial \Sigma_k) \in \mathcal{T}) \leq \begin{cases} C_2 h^2(n) \log(n) & \text{for } d = 2, \\ C_d h^d(n) & \text{for } d \geq 3. \end{cases} \quad (3.9)$$

Since for  $A$  large enough we have  $\mu \geq 3c_d L^d h^d(n) \geq 3c_n$ , Lemma 2.3 yields that for positive constants  $\{c_d, d \geq 2\}$  such that

$$P\left(W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) < \frac{1}{3}\mu\right) \leq \begin{cases} \exp(-\lambda \kappa_2 Ah^2(n) + \lambda^2 c_2 h^2(n) \log(n)) & \text{for } d = 2, \\ \exp(-\lambda \kappa_d Ah^d(n) + \lambda^2 c_d h^d(n)) & \text{for } d \geq 3. \end{cases} \quad (3.10)$$

Thus, after optimizing over  $\lambda$ , we get

$$P\left(\exists \mathcal{T} \subset \Sigma_k : W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) < \frac{1}{3}\mu\right) \leq \begin{cases} n^2 \exp\left(-\frac{\kappa_2^2 A^2 h^2(n)}{4c_2 \log(n)}\right) & \text{for } d = 2, \\ n^d \exp\left(-\frac{\kappa_d^2 A^2 h^d(n)}{4c_d}\right) & \text{for } d \geq 3. \end{cases} \quad (3.11)$$

Thus,  $\{\exists \mathcal{T} \subset \Sigma_k, W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) \leq \frac{1}{3}\mu\}$  has a summable probability if  $A$  is large enough.

### 3.1.3 $\mathcal{C}$ is likely to be covered when $W_{kh(n)}(N \mathbb{I}_0, \mathcal{T})$ is large

We consider here the event  $\{\forall \mathcal{T} \subset \mathcal{S}, W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) \geq \frac{\kappa}{3} Ah^d(n)\}$ . Consider shell  $\mathcal{S}_k$  at a distance  $Ah(n)$  from  $\partial \mathbb{B}(0, n)$ . Since  $\mathcal{S}_k$  is the union of  $\mathbb{B}(z, h(n))$  when  $z \in \Sigma_k$ , Lemma 1.3 implies, when  $d = 2$ , that

$$P(\mathcal{S}_k \not\subset A(N), \forall \mathcal{T}, W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) > \frac{1}{3}\mu) \leq |\mathcal{S}_k| \exp\left(-\kappa_2 \kappa A \frac{h^2(n)}{\log(n)}\right). \quad (3.12)$$

We obtain a summable bound, with  $h(n) = \log(n)$  and  $A$  large enough. When  $d \geq 3$ , then we have

$$P(\mathcal{S}_k \not\subset A(N), \forall \mathcal{T}, W_{kh(n)}(N \mathbb{I}_0, \mathcal{T}) > \frac{1}{3}\mu) \leq |\mathcal{S}_k| \exp(-\kappa_d \kappa Ah^2(n)). \quad (3.13)$$

We obtain a summable bound, with  $h^2(n) = \log(n)$ , and  $A$  large enough.

## 4 Outer error.

In this Section, we prove the outer error estimate (1.5). This is a consequence of our inner error estimates, of Lemma 1.5, combined with coupling with a flashing process of [1]. When dimension  $d = 2$ , and for  $A$  large to be chosen later, we decompose the event  $\{\delta_O(n) \geq A \log(n)\}$ , as

$$\{\delta_O(n) \geq A \log(n)\} = \bigcup_{i \geq 1} \{\delta_O(n) \in [A \log(n) + i - 1, A \log(n) + i]\}. \quad (4.1)$$

In dimension  $d \geq 3$ ,  $\sqrt{\log(n)}$  replaces  $\log(n)$  in (4.1). Note that the index  $i$  is at most  $\log^2(n)$  in view of the results of [1]. Now, we fix  $i \geq 1$ , and we set  $3h(n) = A\sqrt{\log(n)} + i$  in  $d \geq 3$ , and  $3h(n) = A \log(n) + i$  in  $d = 2$ . We consider the event  $\{\delta_O(n) \in [3h(n) - 1, 3h(n)]\}$ . We define also,

$$\Sigma = \mathbb{B}(0, n + 3h(n)) \setminus \mathbb{B}(0, n + 3h(n) - 1).$$

Note now that

$$P(\delta_O(n) \in [3h(n) - 1, 3h(n)]) \leq P\left(\bigcup_{z \in \Sigma} \{z \in A(N), \delta_O(n) = \|z\| - n\}\right). \quad (4.2)$$

For  $z \in \Sigma$ , and in view of Lemma 1.5, we define

$$G(z) = \{z \in A(N), \delta_O(n) = \|z\| - n, |A(N) \cap \mathbb{B}(z, h(n))| > \beta h^d(n)\}. \quad (4.3)$$

We further decompose the right hand side of (4.2) as

$$P(z \in A(N), \delta_O(n) = \|z\| - n) \leq P(G(z)) + P(z \in A(N), |A(N) \cap \mathbb{B}(z, h(n))| \leq \beta h^d(n)). \quad (4.4)$$

The second term on the right hand side of (4.4) is dealt with Lemma 1.5. We deal now with  $G(z)$ . Note that under  $\{\delta_O(n) \in [3h(n) - 1, 3h(n)]\}$ , no explorer escapes  $\mathbb{B}(0, n + 3h(n))$ . Thus, on  $G(z)$ , there are at least  $\beta h^d(n)$  explorers which settle on  $\mathbb{B}(z, h(n))$  before exiting  $\mathbb{B}(0, n + 3h(n))$ . We now express the event  $G(z)$  in term of *flashing explorers*, as introduced in [1].

### 4.1 On a flashing process

We refer the reader to Section 3.1 of [1] for a definition of flashing processes. Here, we partition  $\mathbb{Z}^d$  into shells encaging  $\mathbb{B}(0, n)$ , with for  $k \geq 0$

$$\mathcal{S}_k = \mathbb{B}(0, n + 2(k + 1)h(n)) \setminus \mathbb{B}(0, n + 2kh(n)).$$

Also, for  $k \geq 0$ , let  $\Sigma_k = \partial\mathbb{B}(0, n + (2k + 1)h(n))$ . We now consider the following flashing process. Explorers behave like internal DLA explorers, as long as they stay in  $\mathbb{B}(0, n)$ . After exiting  $\mathbb{B}(0, n)$  they do not flash until their hitting of  $\Sigma_0$ , and behave like *flashing explorers* as defined in Section 3.1 of [1]. In shells  $\{\mathcal{S}_k, k \geq 0\}$ , cells and tiles have the meaning given in Section 4 of [1]. The key features, the reader has to keep in mind, are as follows.

- If a flashing explorer is unsettled up to time  $H(\Sigma_k)$ , then after time  $H(\Sigma_k)$ , it probes one site distributed almost uniformly over the *cell* centered at  $S(H(\Sigma_k))$ , and settles if the site is unoccupied.
- When an explorer leaves the cell centered on  $S(H(\Sigma_k))$ , it cannot afterward settle in  $\mathcal{S}_k$ , but perform a simple random walk, independent of other explorers, until it hits  $\Sigma_{k+1}$ . Thus, if we know that an explorer has reached at time  $t$  a site of  $\mathbb{B}(0, n + (2k + 1)h) \setminus \mathbb{B}(0, n + 2kh)$ , then it performs after time  $t$  a simple random walk, independent of its surrounding, until it reaches  $\Sigma_k$ .
- We can build the internal DLA cluster,  $A(N)$ , and the flashing cluster  $A^*(N)$  using the same trajectories  $S_1, \dots, S_N$  such that

$$A(N) = \bigcup_{i=1}^N \{S_i(T(i))\}, \quad \text{and} \quad A^*(N) = \bigcup_{i=1}^N \{S_i(T^*(i))\}, \quad (4.5)$$

and for all  $i = 1, \dots, N$ ,  $T^*(i) \geq T(i)$ . This last property is fundamental. It implies that if a DLA explorer has crossed a site before settling, then the corresponding flashing explorer has also crossed the site before settling.

Before introducing more notation, let us explain the simple idea behind our estimate.

**Heuristics** Using the representation (4.5), event  $G(z)$  for  $z \in \Sigma$  implies that at least  $\beta h^d(n)$  *flashing explorers* hit  $\mathbb{B}(z, h(n))$  before exiting  $\mathbb{B}(0, n + 3h(n))$ . Consider these explorers after the moment they enter  $\mathbb{B}(z, h(n)) \subset \mathcal{S}_1$  for the first time: they are behaving as independent random walks until they hit  $\Sigma_1$ . Now, a fraction must hit  $\Sigma_1$  on  $\mathbb{B}(z, 2h(n)) \cap \Sigma_1$ . We show that this later event has a probability we can estimate through the approach of [1].

Recall that for  $\Lambda \subset \mathbb{B}(0, n + 3h(n)) \cup \partial\mathbb{B}(0, n + 3h(n))$ , we call  $W_{3h(n)}(N \mathbb{I}_0, \Lambda)$  the number of flashing explorers which cross  $\Lambda$  before hitting  $\Sigma_1$ . Since the initial configuration is always  $N \mathbb{I}_0$  in this section, we omit this coordinate in  $W_{3h(n)}$  to simplify notation. Under our coupling (4.5), we have

$$G(z) \subset \{W_{3h(n)}(\mathbb{B}(z, h(n))) \geq \beta h^d(n)\}. \quad (4.6)$$

Let  $z'$  be the closest site of  $\Sigma_1$  to the line  $(0, z)$ , and note that  $\|z - z'\| \leq 1 + \sqrt{d}$ . Note that a fraction of the  $W_{3h(n)}(B(z, h(n)))$  independent random walks in  $B(z, h(n)) \cap \mathbb{B}(0, n + 3h(n))$ , must hit  $\Sigma_1$  in a neighborhood of  $z'$ . Indeed, first note that since  $z' \in \Sigma_1$ , we have

$$|\partial B(z', 2h(n)) \cap \mathbb{B}(0, n + 3h(n))| \geq \frac{1}{4} |\partial B(z', 2h(n))|. \quad (4.7)$$

Now, for any  $y \in \partial \mathbb{B}(z, h(n))$ , a random walk starting on  $y$ , exits  $\mathbb{B}(z', 2h(n))$  on any site of  $|\partial \mathbb{B}(z', 2h(n))|$  with a probability larger than  $c(2h(n))^{1-d}$ , for some positive constant  $c$ . Thus, there is a positive constant  $\rho$  such that

$$\inf_{y \in \partial B(z, h(n))} P_y(S(H(\partial \mathbb{B}(z', 2h(n)))) \in \mathcal{S}_2) \geq \rho. \quad (4.8)$$

In other words, each flashing explorer stopped on  $\partial \mathbb{B}(z, h(n))$  before hitting  $\Sigma_1$  has a probability at least  $\rho$  to exit  $\Sigma_1$  from  $\Sigma_1 \cap \mathbb{B}(z, 2h(n))$ . Thus, there is a positive constant  $I$ , such that for an integer  $k$  large enough

$$P\left(W_{3h(n)}(\mathbb{B}(z', 2h(n))) < \frac{\rho}{2}k \mid W_{3h(n)}(\mathbb{B}(z, h(n))) > k\right) \leq \exp(-Ik). \quad (4.9)$$

From (4.6), we have

$$\begin{aligned} \bigcup_{z \in \Sigma} G(z) \subset & \bigcup_{z' \in \Sigma_1} \left\{ W_{3h(n)}(\mathbb{B}(z', 2h(n))) \geq \frac{\rho}{2} \beta h^d(n) \right\} \\ & \cup \bigcup_{z' \in \Sigma_1} \left\{ W_{3h(n)}(\mathbb{B}(z', 2h(n))) < \frac{\rho}{2} \beta h^d(n) \text{ and } W_{3h(n)}(B(z, h(n))) \geq \beta h^d(n) \right\}. \end{aligned} \quad (4.10)$$

Let us now define, for any  $a > 0$

$$F(a) = \bigcup_{z \in \Sigma_1} \left\{ W_{3h(n)}(B(z, 2h(n))) \geq ah^d(n) \right\}. \quad (4.11)$$

Thus, from (4.10) and (4.9), and for some constant  $C > 0$

$$\begin{aligned} P\left(\bigcup_{z \in \Sigma} G(z)\right) & \leq P\left(\bigcup_{z' \in \Sigma_1} W_{3h(n)}(\mathbb{B}(z', 2h(n))) \geq \frac{\rho}{2} \beta h^d(n)\right) \\ & \quad + |\Sigma_1| \sup_{z' \in \Sigma_1} P\left(W_{3h(n)}(\mathbb{B}(z', 2h(n))) < \frac{\rho}{2} \beta h^d(n) \mid W_{3h(n)}(B(z, h(n))) \geq \beta h^d(n)\right) \\ & \leq P\left(F\left(\frac{\rho}{2}\beta\right)\right) + Cn^{d-1} \exp\left(-I\frac{\rho}{2}\beta h^d(n)\right). \end{aligned} \quad (4.12)$$

It remains to show that for any fixed  $a$ , we can find  $A$  (defining  $h(n)$ ) such that  $P(F(a))$  is smaller than any inverse power of  $n$ .

## 4.2 On estimating $P(F(a))$ .

Note that by definition of  $\delta_I(n)$ , for  $\mathcal{T} \subset \Sigma_1$   $W_{3h(n)}(\mathcal{T})$  satisfies the inequality

$$W_{3h(n)}(\mathcal{T}) + M_{3h(n)}(\mathbb{B}(0, n - \delta_I(n)), \mathcal{T}) \leq \tilde{M}_{3h(n)}(N \mathbb{I}_0, \mathcal{T}). \quad (4.13)$$

Thus, for  $\alpha_d$  defined in (1.5), we have

$$\mathbb{I}_{\delta_I(n) \leq 2\alpha_d \frac{h(n)}{A}} \left( W_{3h(n)}(\mathcal{T}) + M_{3h(n)}(\mathbb{B}(0, n - 2\alpha_d \frac{h(n)}{A})) \right) \leq \tilde{M}_{3h(n)}(N \mathbb{I}_0, \mathcal{T}). \quad (4.14)$$

Inequality (4.14) puts us in the setting of Lemma 2.4. Thus, we first need to compute

$$\tilde{\mu} = E [M_{3h(n)}(N \mathbb{I}_0, \mathcal{T})] - E \left[ M_{3h(n)}(\mathbb{B}(0, n - 2\alpha_d \frac{h(n)}{A}), \mathcal{T}) \right]. \quad (4.15)$$

Following the same computations as in Section 5.3 of [1], we have for some constants  $K, c > 0$ .

$$\tilde{\mu} \leq K \left( \alpha_d \frac{h(n)}{A} n^{d-1} \right) \times \frac{h^{d-1}(n)}{n^{d-1}} \leq \frac{c}{A} h^d(n). \quad (4.16)$$

Secondly, note that as in Section 5.3 of [1], we have that for constants  $\{c_d, d \geq 2\}$

$$\sum_{z \in \mathbb{B}(0, n)} P_z^2(S(H(\Sigma_1)) \in \mathcal{T}) \leq \begin{cases} c_2 h^2(n) \log(n) & \text{if } d = 2, \\ c_d h^d(n) & \text{if } d \geq 3. \end{cases} \quad (4.17)$$

In optimizing over  $\lambda$  in (2.4), we find for (other) constants  $\{c_d, d \geq 2\}$

$$P(\exists \mathcal{T} : W_{3h(n)}(\mathcal{T}) \geq h^d(n)) \leq P\left(\delta_I(n) > 2\alpha_d \frac{h(n)}{A}\right) + n^d \begin{cases} \exp\left(-c_2 \frac{h^2(n)}{\log(n)}\right) & \text{if } d = 2, \\ \exp\left(-c_d h^d(n)\right) & \text{if } d \geq 3. \end{cases} \quad (4.18)$$

## APPENDIX

### A Proof of Lemma 1.3

We fix  $\eta$ , a configuration of  $AR^d$  explorers in  $\mathbb{B}(0, R/2)$ , and we will choose later  $\alpha$  large enough. Then,

$$P(\mathbb{B}(0, R) \not\subset A(\eta)) \leq \sum_{z \in \mathbb{B}(0, R)} P(W_{\alpha R}(\eta, z) = 0). \quad (A.1)$$

If  $\zeta$  is the configuration with one explorer on each site of  $\mathbb{B}(0, \alpha R) \setminus \mathbb{B}(z, L)$ , we have

$$W_{\alpha R}(\eta, z) + M_{\alpha R}(\zeta, z) \geq \tilde{M}_{\alpha R}(\eta, z) - |\mathbb{B}(z, L)|. \quad (A.2)$$

Note that  $W_{\alpha R}(\eta, z)$  and  $M_{\alpha R}(\zeta, z)$  are independent, so that we are in the setting of Lemma 2.3. Assume for a moment that conditions (H1) and (H2) hold, and in addition

$$E[M_{\alpha R}(\eta, z)] - E[M_{\alpha R}(\zeta, z)] \geq \max \left( 3|\mathbb{B}(z, L)|, \sum_{y \in \mathbb{B}(0, \alpha R)} P_y(H_z < H_{\alpha R})^2 \right). \quad (\text{A.3})$$

Then, we have

$$P(W_{\alpha R}(\eta, z) = 0) \leq \exp(-C(E[M_{\alpha R}(\eta, z)] - E[M_{\alpha R}(\zeta, z)])). \quad (\text{A.4})$$

It would remain to show that for some constants  $c_d$

$$E[M_{\alpha R}(\eta, z)] - E[M_{\alpha R}(\zeta, z)] \geq \begin{cases} c_2 \frac{AR^2}{\log(R)} & \text{if } d = 2, \\ c_d AR^2 & \text{if } d \geq 3. \end{cases} \quad (\text{A.5})$$

For this purpose, we next consider separately the case  $d \geq 3$  and the case  $d = 2$ , and show (A.5) and (A.3).

## A.1 The case $d \geq 3$

We show in this Section that for  $\kappa_d > 0$ , and  $A$  large enough,

$$E[\tilde{M}_{\alpha R}(\eta, z) - M_{\alpha R}(\zeta, z)] \geq \frac{\kappa}{2} AR^2 \gg 3|\mathbb{B}(z, L)|. \quad (\text{A.6})$$

The proof is based on the following classical estimates. There is  $a_1, a_2$  positive constants such that for any  $y, z \in \mathbb{Z}^d$

$$\frac{a_1}{1 + \|y - z\|^{d-2}} \leq P_y(H_z < \infty) \leq \frac{a_2}{1 + \|y - z\|^{d-2}}. \quad (\text{A.7})$$

Note first that when  $L$  is large enough, (H1) holds. Indeed,

$$\sup_{y: \|z-y\| > L} P_y(H_z < H_{\alpha R}) \leq \sup_{y: \|z-y\| > L} P_y(H_z < \infty) \leq \frac{a_2}{1 + L^{d-2}} \leq \frac{\kappa - 1}{\kappa}, \quad \text{with } \kappa > 1. \quad (\text{A.8})$$

We now estimate the mean number of explorers hitting  $z$ .

$$\begin{aligned} E[M_{\alpha R}(\eta, z)] - E[M_{\alpha R}(\zeta, z)] &= \sum_{y \in \mathbb{B}(0, R/2)} \eta(y) P_y(H_z < H_{\alpha R}) - \sum_{y \in \mathbb{B}(0, \alpha R) \setminus \mathbb{B}(z, L)} P_y(H_z < H_{\alpha R}) \\ &\geq \sum_{y \in \mathbb{B}(0, R/2)} \eta(y) P_y(H_z < H_{\alpha R}) - \sum_{y \in \mathbb{B}(0, \alpha R)} P_y(H_z < \infty). \end{aligned} \quad (\text{A.9})$$

Note that for  $y \in \mathbb{B}(0, R/2)$ , we have

$$\begin{aligned} P_y(H_z < H_{\alpha R}) &= P_y(H_z < \infty) - E_y \left[ \mathbb{1}_{H_{\alpha R} < H_z} P_{S(H_{\alpha R})}(H_z < \infty) \right] \\ &\geq \frac{a_1}{1 + \|y - z\|^{d-2}} - E_y \left[ \frac{a_2}{1 + \|S(H_{\alpha R}) - z\|^{d-2}} \right] \\ &\geq \inf_{y \in \mathbb{B}(0, R/2)} \frac{a_1}{1 + \|y - z\|^{d-2}} - \sup_{y \in \partial \mathbb{B}(0, \alpha R)} \frac{a_2}{1 + \|y - z\|^{d-2}}. \end{aligned} \quad (\text{A.10})$$

Now, for a constant  $\alpha$  which depends only on  $a_1, a_2$ , there is  $\kappa > 0$  such that

$$\inf_{y \in \mathbb{B}(0, R/2)} P_y(H_z < H_{\alpha R}) \geq \frac{\kappa}{R^{d-2}}. \quad (\text{A.11})$$

Now, using (A.11) in (A.9), we have a constant  $c$  such that

$$\begin{aligned} E[\tilde{M}_{\alpha R}(\eta, z) - M_{\alpha R}(\zeta, z)] &\geq AR^d \frac{\kappa}{R^{d-2}} - \sum_{y: \|y-z\| < \alpha R} \frac{a_1}{1 + \|y-z\|^{d-2}} \\ &\geq \kappa AR^2 - ca_2(\alpha R)^2. \end{aligned} \quad (\text{A.12})$$

When  $A$  is chosen large enough, we obtain (A.6).

Finally, there are constants  $\{C_d, d \geq 3\}$  such that for any  $z \in \mathbb{B}(0, R)$

$$\sum_{y \in \mathbb{B}(0, \alpha R)} P_y(H_z < H_{\alpha R})^2 \leq \begin{cases} C_3 \alpha R & \text{for } d = 3, \\ C_4 \log(\alpha R) & \text{for } d = 4, \\ C_d & \text{for } d \geq 5. \end{cases} \quad (\text{A.13})$$

Thus, hypotheses (A.3) holds.

## A.2 The case $d = 2$

We still have

$$P_y(H_z < H_{\alpha R}) = \frac{G_{\alpha R}(y, z)}{G_{\alpha R}(z, z)} = \frac{G_{\alpha R}(z, y)}{G_{\alpha R}(z, z)}, \quad \text{and} \quad G_{\alpha R}(z, y) = E_z[a(S(H_{\alpha R}), y)] - a(z, y), \quad (\text{A.14})$$

where the *potential kernel*  $a(\cdot, \cdot)$  replaces Green's function. Note that for  $0 \leq \|z\| + R < \alpha R$ , we have two positive constants  $K_2$  and  $K'_2$  such that

$$K'_2 \log(2\alpha R) \geq G_{B(z, 2\alpha R)} \geq G_{\alpha R}(z, z) \geq G_{B(z, R)}(z, z) \geq K_2 \log(R), \quad (\text{A.15})$$

by Proposition 1.6.6 of Lawler [4]. To estimate  $G_{\alpha R}(z, y)$ , we use Theorem 4.4.4. of [6] which establishes that for  $z \neq 0$ , (with  $\gamma$  the Euler constant)

$$\left| a(0, z) - \frac{2}{\pi} \log(\|z\|) - \frac{2\gamma + \log(8)}{\pi} \right| \leq \frac{K_g}{\|z\|^2}. \quad (\text{A.16})$$

Thus, for  $y \in \mathbb{B}(0, \alpha R)$ , and  $0 \leq \|z\| \leq R$ , and  $y \neq z$

$$\left| G_{\alpha R}(z, y) - \frac{2}{\pi} E \left[ \log \left( \frac{\|S(H_{\alpha R}) - z\|}{\|y - z\|} \right) \right] \right| \leq 2K_g. \quad (\text{A.17})$$

When  $y \in B(0, R/2)$  we get

$$G_{\alpha R}(z, y) \geq \frac{2}{\pi} \log(2(\alpha - 1)/3) - 2K_g. \quad (\text{A.18})$$

We choose  $\alpha$  large enough so that for some constant  $C_1$ , we have, for all  $y$  in  $B(0, R/2)$ ,

$$G_{\alpha R}(z, y) \geq C_1. \quad (\text{A.19})$$

Formulas (A.14), (A.15) and (A.19) together imply that

$$E[M_{\alpha R}(\eta, z)] = \sum_{y \in \mathbb{B}(0, R/2)} \eta(y) P_y(H_z < H_{\alpha R}) \geq \frac{C_1}{K'_2 \log(2\alpha R)} \sum_{y \in \mathbb{B}(0, R/2)} \eta(y) = \frac{C_1 A R^2}{K'_2 \log(2\alpha R)}. \quad (\text{A.20})$$

Using Lemma 3 of [5], we have, for some positive constant  $C_2$ ,

$$E[M_{\alpha R}(\zeta, z)] \leq E[M_{\alpha R}(\mathbb{B}(0, \alpha R), z)] \leq \frac{C_2 (\alpha R)^2}{\log(R)}. \quad (\text{A.21})$$

We need now to choose  $L$  to have (H1) satisfied. Note that for  $y \neq z$ , (A.17) and (A.15) yields

$$P_y(H_z < H_{\alpha R}) \leq \frac{1}{K_2 \log(R)} E \left[ \frac{2}{\pi} \log \left( \frac{\|S(H_{\alpha R}) - z\|}{\|y - z\|} \right) + 2K_g \right]. \quad (\text{A.22})$$

If  $\|z - y\| > R/\log(R)$ , we obtain, for some constant  $C_3$

$$P_y(H_z < H_{\alpha R}) \leq \frac{C_3 \log((\alpha + 1) \log(R))}{\log(R)}. \quad (\text{A.23})$$

When  $R$  is large enough, we have that (H1) holds for  $L = R/\log(R)$ . Note that  $|\mathbb{B}(0, L)|$  is of order  $R^2/\log(R)^2$  and is much smaller than  $R^2/\log(R)$ .

Finally we need to control the sum of second moments. Simply note that, from (A.21),

$$\sum_{y \in \mathbb{B}(0, \alpha R) \setminus \mathbb{B}(0, L)} P_y^2(H_z < H_{\alpha R}) \leq E[M_{\alpha R}(\zeta, z)] \leq \frac{C_2 \alpha^2 R^2}{\log(R)}. \quad (\text{A.24})$$

## B Proof of Lemma 1.6

We will choose an  $h$  such that  $R/2h$  is a positive integer. We divide  $\mathcal{S} = B(0, 2R) \setminus B(0, R)$  into  $R/2h$  concentric shells of height  $2h$ . For  $i = 1, \dots, R/2h$  define

$$\mathcal{S}_k = \mathbb{B}(0, 2R - 2(k-1)h) \setminus \mathbb{B}(0, 2R - 2kh), \quad \text{and} \quad \Sigma_k := \partial \mathbb{B}(0, 2R - (2k-1)h). \quad (\text{B.1})$$

Also, we set  $\mathcal{S}_0 = \mathbb{B}(0, 2R)^c$ . Then, we start on  $z \in \partial \mathbb{B}(0, 2R)$  a flashing explorer associated with this partition with an explored region  $V$ . The flashing setting is much simpler than the one introduced in Section 3.1 of [1]. There is an underlying simple random walk, say  $S^*$ , and each shell  $\mathcal{S}_1, \mathcal{S}_2, \dots$  is associated with a flashing site. These flashing sites, say  $\{Z_k, 0 \leq k \leq 2R/h\}$  are obtained as follows. For  $k = 0$  we set  $Z_0 = z$  and for  $k \geq 1$  we draw a continuous random variable  $R_k$  on  $[0, h]$  with density in  $r \in [0, h] \mapsto dr^{d-1}/h^d$ : the flashing site  $Z_k$  is the exit site from  $\mathbb{B}(S^*(H(\Sigma_k)), R_k)$  after time  $H(\Sigma_k)$ . Then, the explorer settles on the first flashing site in  $\mathcal{S} \setminus V$ . The purpose of the flashing construction is that (i) the flashing site is distributed almost uniformly inside the ball  $\mathbb{B}(S^*(H(\Sigma_k)), h)$ , and (ii)  $P_z(H(\mathbb{B}(0, R)) < H(V^c))$  is bounded above by the probability that the explorer crosses  $\mathcal{S}$ .

For a small  $\beta$  to be chosen later, we say that  $y \in \Sigma_k$  has a *dense neighborhood* if  $|\mathbb{B}(y, h) \cap V| \geq \beta h^d$ , and we call  $D_k$  their set. There is  $\kappa > 0$  such that for  $h$  large enough (say  $h \geq h_0$ )

- if  $S^*(H(\Sigma_k)) \notin D_k$ , then the probability that  $S^*$  does not settle in  $\mathcal{S}_k$  is smaller than  $\kappa\beta$ ,
- the probability that  $S^*(H(\Sigma_k)) \in D_k$  is smaller than  $\kappa|D_k|/h_k^{d-1}$  (see Lemma 5 of [3]) uniformly over the of position the previous flashing site (in  $\mathcal{S}_{k-1}$  or, exceptionally, on the border of  $\mathcal{S}_{k-1}$ ).

Now, the flashing explorer has crossed the annulus  $\mathcal{S}$  if  $Z_k \in V$  for all  $k \geq 1$ . In other words,

$$\{H(\mathbb{B}(0, R)) < H(V^c)\} \subset \bigcap_{k=1}^{R/2h} \{Z_k \in V\}. \quad (\text{B.2})$$

By conditioning, we obtain

$$P_z \left( \bigcap_{k=1}^{R/2h} \{Z_k \in V\} \right) \leq \prod_{k=1}^{R/2h} \left( \kappa\beta + \frac{\kappa|D_k|}{h_k^{d-1}} \right). \quad (\text{B.3})$$

By the arithmetic-geometric inequality, and (B.2), we obtain

$$P_z(H(\mathbb{B}(0, R)) < H(V^c)) \leq \left( \kappa\beta + \frac{\kappa}{R/2h} \sum_{k=1}^{R/2h} \frac{|D_k|}{h^{d-1}} \right)^{R/2h}. \quad (\text{B.4})$$

Note that each  $y \in D_k$  satisfies  $|\mathbb{B}(y, h) \cap V| \geq \beta h^d$ , but each site in  $\mathbb{B}(y, h) \cap V$  is in the neighborhood of at most  $h^{d-1}$  site of  $D_k$ . Thus, for some  $\kappa'$

$$\sum_{k=1}^{R/2h} \frac{\beta|D_k|h^d}{h^{d-1}} \leq \kappa'|V|, \quad \text{i.e.,} \quad \frac{1}{R/2h} \sum_{k=1}^{R/2h} \frac{|D_k|}{h^{d-1}} \leq \frac{2\kappa'|V|}{\beta R h^{d-1}}. \quad (\text{B.5})$$

We choose now  $\beta$  such that  $4\kappa\beta < 1$ , and we choose the smallest  $h$  such that  $R/2h$  is a positive integer and

$$h \geq \max \left\{ h_0, \left( \frac{2\kappa'|V|}{\beta^2 R} \right)^{\frac{1}{d-1}} \right\}. \quad (\text{B.6})$$

This adds a constraint on  $|V|$ :

$$|V| \leq \frac{\beta^2}{2^d \kappa'} R^d. \quad (\text{B.7})$$

Instead of including (B.7) as a condition of our Lemma, we find more convenient to note that the probability we estimate is always less than 1, so that we deal with the case where (B.7) is violated with the constant  $a_d$  of (1.11).

## C Proof of Lemma 1.5

Recall that  $a_d$  and  $\kappa_d$  are the constants appearing in Lemma 1.6. We define a positive constant

$$\gamma = \max \left( 1, \left( \frac{2a_d}{\kappa_d} \right)^{d-1} \right). \quad (\text{C.1})$$

Choose now  $\beta > 0$  such that  $4^d \beta \gamma \leq 1$ , and  $h_0 = R/4 \geq 1$ . Note that

$$\gamma|\eta| \leq \gamma\beta R^d \leq h_0^d. \quad (\text{C.2})$$

We build now, by induction, a random subdivision of  $\mathbb{B}(z, R)$  into shells of heights  $h_0, h_1, \dots$ , in which respectively  $N_0, N_1, \dots$  explorers of  $A(\eta)$  have settled. We emphasize that the randomness comes from  $A(\eta)$ , and that the event  $\{0 \in A(\eta)\}$  imposes to have  $N_i \geq \lfloor h_i \rfloor$ , for  $i \geq 0$ . Assume that  $h_1, \dots, h_k$  have been defined such that

$$h_k \geq 1 \quad \text{and} \quad \sum_{i=1}^k h_i < \frac{R}{2}. \quad (\text{C.3})$$

We define  $h_{k+1}^d = \gamma N_k \leq \gamma|\eta|$ , and, by (C.2) we have  $h_{k+1} \leq h_0$ . Note also that  $h_{k+1} \geq 1$ . Indeed, necessarily  $N_k \geq \lfloor h_k \rfloor$ , so that  $h_{k+1}^d \geq \gamma \lfloor h_k \rfloor \geq \lfloor h_k \rfloor$ . Since  $\min(h_1, \dots, h_{k+1}) \geq 1$ , the number of steps before we violate (C.3), say  $L$ , is finite. Obviously  $L \leq R$ . Note that since  $h_L \leq h_0$

$$\frac{R}{2} \leq \sum_{i=1}^L h_i \leq h_L + \sum_{i=1}^{L-1} h_i \leq \frac{R}{4} + \frac{R}{2}. \quad (\text{C.4})$$

Thus, we define

$$h_{L+1} = R - \left( \sum_{i=0}^L h_i \right) \geq 0. \quad (\text{C.5})$$

For any choice of integers  $l, n_0, \dots, n_l$ , the event  $\{L = l, N_0 = n_0, \dots, N_L = n_l\}$  implies that  $n_1 + \dots + n_l$  explorers have crossed a shell  $B(z, R) \setminus B(z, R - h_0)$  by stepping on at most  $n_0$  explorers settled in it, that  $n_2 + \dots + n_l$  explorers have crossed shell  $B(z, R - h_0) \setminus B(z, R - h_0 - h_1)$  with  $n_1$  explorers settled in it, and so on and so forth. Using Lemma 1.6, the fact that  $n_i \leq \beta R^d$ ,  $l \leq R$ , and the notation  $\delta = \frac{1}{d-1}$ , we reach the following estimate.

$$\begin{aligned} & P(0 \in A(\eta)) \\ & \leq \sum_{\substack{l \leq R, n_0, n_1, \dots, n_l \leq |\eta| \\ \forall i, n_i \geq \lfloor h_i \rfloor}} P(L = l, N_0 = n_0, \dots, N_L = n_l) \\ & \leq R(\beta R^d)^{R+1} \sup_{\substack{l \leq R, n_0, n_1, \dots, n_l \leq |\eta| \\ \forall i, n_i \geq \lfloor h_i \rfloor}} e^{a_d \sum_{i=1}^L i n_i} \exp \left( -\kappa_d \sum_{i=1}^L n_i \left( \left( \frac{h_0^d}{n_0} \right)^\delta + \dots + \left( \frac{h_{i-1}^d}{n_{i-1}} \right)^\delta \right) \right). \end{aligned} \quad (\text{C.6})$$

Now, note that by the arithmetic-geometric inequality, for  $1 \leq i \leq l$  (and using  $h_i \leq h_0$ )

$$\begin{aligned} \frac{1}{i} \left( \left( \frac{h_0^d}{n_0} \right)^\delta + \dots + \left( \frac{h_{i-1}^d}{n_{i-1}} \right)^\delta \right) & \geq \left( \frac{h_0^d}{n_0} \times \dots \times \frac{h_{i-1}^d}{n_{i-1}} \right)^{\delta/i} \\ & = \left( \frac{h_0^d}{n_{i-1}} \gamma^{i-1} \right)^{\delta/i} = \left( \frac{h_0^d}{h_i^d} \gamma^i \right)^{\delta/i} \geq \frac{2a_d}{\kappa_d}. \end{aligned} \quad (\text{C.7})$$

Thus, from (C.6) and (C.7), we have

$$\begin{aligned}
 P(0 \in A(\eta)) &\leq R(\beta R^d)^{R+1} \max_{\substack{l \leq R, n_0, n_1, \dots, n_l \leq |\eta| \\ \forall i, n_i \geq \lfloor h_i \rfloor}} \exp\left(-a_d \sum_{i=1}^L i n_i\right) \\
 &\leq R(\beta R^d)^{R+1} \max_{\substack{l \leq R, n_0, n_1, \dots, n_l \leq |\eta| \\ \forall i, n_i \geq \lfloor h_i \rfloor}} \exp\left(-\frac{a_d}{\gamma} \sum_{i=1}^{L-1} i h_{i+1}^d\right)
 \end{aligned} \tag{C.8}$$

Since  $h_1 \leq R/4$ , note that we have  $h_2 + \dots + h_L \geq R/4$  by (C.4). By Hölder inequality, note that for constants  $\{c_d, d \geq 2\}$

$$\sum_{i=1}^{L-1} i h_{i+1}^d \geq \frac{\left(\sum_{i=1}^{L-1} h_{i+1}\right)^d}{\left(\sum_{i=1}^{L-1} \frac{1}{i^{1/(d-1)}}\right)^{d-1}} \geq \begin{cases} c_2 \frac{R^2}{\log(L)} \geq c_2 \frac{R^2}{\log(R)} & \text{for } d = 2, \\ c_d \frac{R^d}{L^{d-2}} \geq c_d R^2 & \text{for } d \geq 3. \end{cases} \tag{C.9}$$

This concludes the proof.

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