Games and Discrimination Lessons From *The Weakest Link*

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ABSTRACT

We use data from the television game show, The Weakest Link, to determine whether contestants discriminate on the basis of race and gender and, if so, which theory of discrimination best explains their behavior. Our results suggest no evidence of discriminatory voting patterns by males against females or by whites against blacks. In contrast, we find that in the early rounds of the game women appear to discriminate against men. We test three theories for the voting behavior of women: preference-based discrimination, statistical discrimination, and strategic discrimination. We find only preferencebased discrimination to be consistent with the observed voting patterns.

I. Introduction

The earnings of both women and blacks have consistently lagged behind those of white men. Empirically determining whether these differences arise because of discrimination is extremely difficult, and distinguishing between the various theories of discrimination is harder still. As a result, researchers have begun to look outside of labor markets to develop a better understanding of whether and why individuals discriminate. This paper builds on the emerging experimental literature on discrimination and explores the use of a high-stakes game environment to reveal patterns of discrimination. To this end, we use data from the television game show *The*

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Weakest Link to determine whether contestants discriminate on the basis of race and gender, and, if so, which theory of discrimination best explains their behavior.

A number of existing papers analyze contestant behavior on television game shows in order to draw inferences about behavior in more policy-relevant settings.¹ The benefit of this type of analysis is twofold. First, in many situations, the strategic environment of game shows generates data that can be used to answer questions that cannot be easily addressed with data from more traditional settings. This is particularly relevant in the context of labor market discrimination where economists have had only limited success at distinguishing between the various theories of discrimination using standard data. Second, the stakes are much larger on game shows than in laboratory experiments and these stakes may influence both the extent and nature of discrimination. For example, Ball and Eckel (2003) find that favoritism toward members of "high status" groups in ultimatum games is sensitive to the dollar value of the prize being split.

The principal feature of The Weakest Link that enables us to study discrimination is the fact that players have multiple opportunities to cast votes against their fellow contestants in order to remove them from the game. As a result, we are able to examine whether race and gender play a role in determining voting patters. In addition, we are aided by the fact that the data from the show provide excellent controls for individual ability. This is important since one of the principal difficulties in establishing the presence of discrimination is that it is almost impossible to determine whether unequal outcomes arise because of discrimination or because of unobservable race and gender differences in productivity. However, in The Weakest Link we observe the same explicit measure of individual ability-the percentage of questions answered correctly-that is observed by the contestants. Our ability to observe a large fraction of the information that the contestants have about one another limits the extent of the omitted variables bias problem that plagues discrimination research. Further, even if there remain important unobserved characteristics, the structure of the game allows us to control for these omitted factors by using information on the voting patterns of the other contestants. In particular, if we are trying to examine the factors that determine whether Player A votes against Player B, we can control for the proportion of other contestants who vote against B. In this way, we can account for other factors that are not directly included in our data but that are relevant to the players' voting decisions. In addition, because data are collected from two slightly different versions of the show, a daily show and a weekly show, the veracity of our results can be tested against two independent samples.

The structure of the game also allows us to examine why contestants discriminate. In particular, we are able to distinguish preference-based discrimination (in which discrimination arises because people simply do not like members of certain groups) from statistical discrimination (in which discrimination arises because people use group identity as a proxy for unobserved ability). In a strategic environment such as *The Weakest Link*, we also need to consider a third type of discrimination: what we term "strategic discrimination." The notion here is that in a strategic setting race and gender may serve as focal points for collusive behavior (see, for example, Holm 2000).

^{1.} For example, Bennett and Hickman (1993), Berk, Hughson, and Vandezande (1996), Gaines (2003), Gertner (1993), Gilbert and Hatcher (1994), Metrick (1995), and Tenorio and Cason (2002).

A major contribution of this paper is to highlight various methods of distinguishing between these three theories of racial discrimination.

To look for evidence of discriminatory behavior, we evaluate the voting decisions of individual players using a series of conditional logit models. A major advantage of the conditional logit model relative to analysis of the total votes cast against a particular group is that it allows us to consider interactions between the characteristics of the individuals doing the voting and the characteristics of those receiving the votes. Thus, for example, one can examine whether men are more likely to vote against women *and* whether women are more likely to vote against men. The results reveal no evidence of discrimination against either women or blacks. However, we consistently find that women are substantially more likely to vote against men in the early rounds of the game, even after controlling for a broad set of performance measures. That is, based on these findings, it appears that women discriminate against men (and in favor of women) in their voting decisions. We are able to show that neither statistical discrimination nor strategic discrimination explains this pattern. Instead, the evidence seems broadly consistent with women simply preferring to play the game with other women.

At first glance, it may seem surprising that we find no evidence of discrimination against women or blacks. However, the data reveal no significant race or gender differences in contestants' ability to answer questions correctly. As a result, even if statistical discrimination is a feature of the labor market, we are able to rule it out as a possible source of discrimination in our data.² Thus, given that we find no evidence of preference-based discrimination or strategic discrimination against women or blacks, our results point toward statistical discrimination as a possible explanation for the discriminatory patterns found in studies of the labor market.

Interestingly, the notion that women have discriminatory preferences in favor of other women is reflected in a number of recent studies. For example, Dillingham, Ferber, and Hamermesh (1994) examine male and female voting behavior in the selection of officers for a professional association. Confirming our results, they find that men do not discriminate on the basis of gender while women discriminate in favor of female candidates.

II. Previous Work on Racial Discrimination

A. Evidence From the Labor Market

It is not surprising that there exists an enormous literature on racial and gender discrimination in the labor market. Here we divide these studies into two groups: those that test for evidence of discrimination and those that attempt to distinguish between different theories of discrimination.

One approach to testing for evidence of discrimination is to run wage regressions that control for individual-level characteristics (see, for example, Altonji and Blank 1999 for a full discussion). However, it is well-known that the estimated race and gender wage

^{2.} A possible exception would be the case where contestants have incorrect prior beliefs about one another's ability. We test for this possibility and find no evidence to support for the presence of incorrect priors.

differentials from these regressions should not necessarily be interpreted as evidence of discrimination. The primary problem is that unobservable individual-level characteristics that are correlated with race and gender may lead to estimated race and gender wage differentials even in the absence of discrimination. A second problem is endogeneity bias. Many individual-level characteristics that we would like to include as controls may be influenced by discrimination. For example, because blacks tend to have lower levels of experience than whites, it would seem natural to include experience in the wage regressions. However, if blacks have lower levels of experience because of past discrimination, then including experience will lead the coefficient on race to understate the true extent of discrimination. Thus, while wage regressions often provide interesting descriptions of the data, they do not provide solid evidence of discrimination.

A second strand of the literature uses audit studies to look for evidence of discrimination (Cross, Kenny, and Zimmerman 1990; Turner, Fix, and Struyk 1991; Neumark 1996). The majority of these studies find substantial evidence of discrimination against women and blacks in both the labor and housing markets. However, as Heckman and Siegelman (1992) and Heckman (1998) point out, these studies may fail to overcome the omitted variables bias that plagues standard regression techniques. In particular, the implicit assumption of audit studies is that the auditors are identical apart from their race and gender. However, the researcher who selects the auditors may not know or may not be able to observe all of the characteristics that are relevant to employers. Thus, if there is any relationship between race and gender and these unobserved components, then the differences in outcomes across these groups may not truly represent discrimination.

Two studies that circumvent this problem are Goldin and Rouse (2000) and Bertrand and Mullanaithan (2004). Goldin and Rouse examine whether the adoption of "screens" that hide the identity of auditioning musicians from judges has been responsible for the increase in the proportion of females hired at orchestras. Overall their evidence is consistent with discrimination against female musicians, and indicates that the use of the screen has been responsible for a large fraction of the increase in the number of women hired. Bertrand and Mullanaithan (2004) conduct a field experiment in which they send out a series of resumes to employers in the Boston and Chicago areas that are essentially identical apart from the name at the top of the page.³ They find that the response rates on resumes with white-sounding names are 50 percent higher than those with African American names. While both of these provocative studies find evidence of discrimination, they do not fully reveal *why* the discrimination occurs.

Within the field of economics, the two leading theories of discrimination are preference-based discrimination and statistical discrimination. In models of preferencebased discrimination, employers act as if there is some cost associated with hiring workers from a particular group, and, in equilibrium, workers from these groups are paid lower wages than workers from other groups (Becker 1957). In contrast, in models of statistical discrimination firms cannot perfectly observe worker productivity, and base their assessment of worker productivity on prior beliefs about the productivity of workers in different groups. Thus, discrimination can arise either because of group differences in average productivity or in the quality of information that firms

^{3.} Neumark, Bank, and Van Nort (1995) also partially address this critique by emphasizing the use of resumes in the preinterview stage.

have about workers. (See, for example, Phelps 1972; Arrow 1973; Aigner and Cain 1977; Lundberg and Startz 1983; Coate and Loury 1993.)

A handful of studies explicitly test for evidence of preference-based discrimination. Three of these, Ashenfelter and Hannan (1986), Hellerstein, Neumark, and Troske (2002), and Kahn and Sherer (1988), find evidence consistent with preference-based discrimination. A smaller number of studies test for evidence of statistical discrimination. For example, both Oettinger (1996) and Altonji and Pierret (2001) build models of statistical discrimination in which employers slowly learn about worker quality. Unfortunately, their models generate conflicting predictions about what differences in black-white age-earnings profiles imply about statistical discrimination.⁴

B. Evidence From Experiments

Complementing the empirical work on labor markets are a number of papers by both psychologists and economists that look for evidence of discrimination in laboratory settings. A common finding is that individuals display "in-group bias" in the sense that they favor members of their own group (for example, Turner 1978; Turner and Brown 1978; Vaughn, Tajfel, and Williams 1981). The evidence also suggests that group status plays a role in determining the outcome of bargaining games (Ball et al. 2001). A number of papers also look for evidence of statistical discrimination in laboratory settings (for example, Anderson and Haupert 1999; Davis 1987; Fryer, Goeree, and Holt 2005). These papers consistently show that statistical discrimination tends to arise in the presence of incomplete information. Generally, the groups defined in these settings are artificial. (For example, the groups might be "green" and "yellow.") One exception is Fershtman and Gneezy (2001) who analyze discrimination among Ashkenazic and Eastern Jews. They find discrimination against Eastern Jews in "trust" games by both Ashkenazic and Eastern Jews. Conducting a series of experiments, Fershtman and Gneezy argue that this result is due to incorrect expectations regarding the "trustworthiness" of Eastern Jews.

Simultaneous to our paper, Levitt (2004) also examines discrimination in *The Weakest Link.*⁵ Our approaches are very different. Levitt focuses on aggregate discrimination (total votes cast against a particular group) rather than whether discrimination against a given group depends on the demographics of those doing the voting. For instance, we are interested in determining whether women vote against men and whether men vote against women, effects that are not observable and may cancel each other out under Levitt's approach.⁶ Levitt's results also differ from our own because of differences in our interpretation of the strategic incentives in the game. In particular, Levitt's method relies on two assumptions. First, in Levitt's paper the extent of

^{4.} Oettinger interprets the earnings gap between blacks and whites increasing over the life-cycle as evidence for statistical discrimination while Altonji and Pierret interpret the same feature as evidence against statistical discrimination.

^{5.} Indeed, in late December of 2002, we both found out the other was close to completing their project. The result was two January 2003 working papers.

^{6.} Additionally, because we use a conditional logit, the characteristics of the other contestants affect the probability of receiving a vote and the total number of predicted votes will always equal the number of votes cast. Models that simply regress the number of votes cast against on one's characteristics will predict more votes than are actually available for shows with weak contestants.

taste-based discrimination must be the same across all rounds of the game. However, because the implicit cost of taste-based discrimination rises as the game progresses (because one's probability of winning the game is higher in later rounds), discriminatory outcomes due to taste discrimination should diminish over time. Second, Levitt assumes that voting incentives switch as the game progresses, so that contestants first want to vote off weak players and later want to vote off strong players. Given the structure of the game, it is not clear that the voting incentives will truly reverse. We find that although the incentives to vote off the weakest player diminish as the game progresses, at no point is the strongest player ever more likely to be voted off than the weakest player. An advantage of Levitt's work is that he has a larger sample of daily shows and analyzes discrimination against the elderly and Hispanics.

III. Data and Rules of the Game

We focus on the version of *The Weakest Link* produced by NBC Enterprises and broadcast in the United States. Our data come from our own video recordings of more than 100 episodes.

There are two versions of *The Weakest Link* an hour long weekly show and a halfhour long daily show, with both versions following the same general structure. After excluding celebrity episodes where the contestants play for charity, our data consist of 28 weekly shows and 75 daily shows.⁷ Each show is divided into a series of timed rounds, with the number of rounds corresponding to the number of players: eight rounds in the weekly show and six rounds in the daily show. Within each round, players are sequentially asked to answer general trivia questions where correct answers translate to an increase in the prize money. The first correct answer is worth \$1,000 in the weekly show and \$250 in the daily show. After a correct answer, a player can choose to "bank" the money for the team. If the player banks, the next correct answer is again worth \$1,000 in the weekly show and \$250 in the daily show. Should the player decide not to "bank," the amount of money added to the pot following a correct answer increases. However, failure to answer a question correctly leads to the loss of any unbanked money for that round. A successive chain of eight (six) correct answers with no intermittent "banks" leads to a \$125,000 (\$12,500) increase in the pot. Money banked from each round is accumulated into a team bank.

After each round, each player votes independently as to which player he would like to remove from the show, and the player who receives the most votes must leave the game. In the event of a tie, the "strongest link" chooses which player to remove from the subset of players who received the most votes. The strongest link is the player who answers the highest percentage of his or her questions correctly. Once the field of players is reduced to two (this occurs in Round 7 of the weekly show and Round 5 of the daily show), these two players first accumulate prize money in the same fashion as in the earlier rounds, after which the two players compete directly against each other with the winner taking all the money in the team bank.⁸

^{7.} The data are incomplete for some rounds due to broadcast interruptions.

^{8.} In the final round, five (three) questions are asked of each contestant in the weekly (daily) discrimination by women against men in the early rounds of the game.

One drawback of our data is that the contestants on *The Weakest Link* are unlikely to be representative of the U.S. population. Indeed, this problem plagues much of the experimental literature. We have investigated how contestants are selected to be on the *The Weakest Link*. The show draws contestants from a diverse set of groups. Besides variation in gender and race, contestants come from a broad range of ages, occupations and educational backgrounds. As with most television game shows, the most significant criterion for selection is whether the contestant is telegenic. We have no a priori reason to believe this is correlated with discriminatory behavior.

IV. Why Discriminate on The Weakest Link?

The Weakest Link provides a unique environment in which to distinguish between various theories of discrimination. Playing the game well not only involves answering questions correctly, but also making astute inferences about the other players' ability to play the game. Because we observe players' decisions about who to vote off, we can determine whether race and gender are relevant factors in voting decisions even after controlling for each player's performance.

In addition, we examine why contestants discriminate. Complicating this analysis is the game's evolving strategic environment. In the early stages of the game, when the contestants act cooperatively to build the pot of prize money, there is a clear incentive to vote off weak players. Indeed, in Round 1 of the daily show, the weakest player is voted off over 50 percent of the time, while the strongest player is voted off only 4 percent of the time. However, voting incentives may shift as the game progresses. As discussed above, in the final round, the two remaining players first build the pot of prize money cooperatively as they do in earlier rounds. Then they face one another in a head-to-head competition to determine the winner of the entire game. Thus, concern about the head-to-head competition creates an incentive to vote off strong contestants, particularly in later rounds. It is unclear, however, whether the incentive to vote off strong players will outweigh the incentive to vote off the weak players because even in the final round the players cooperatively build the prize money. Empirically, we find that even though the probability of voting off the weakest link diminishes as the game progresses, players are always more likely to vote off the weakest link than the strongest link. Nonetheless, due to the ambiguity in voting incentives and the complications associated with updating beliefs and incorporating past voting behavior into current voting behavior, we base the majority of our analysis on the first round of the game where voting incentives are more clear-cut.

The remainder of this section presents a brief discussion of the expected voting behavior under the assumptions of (a) statistical discrimination, (b) preference-based discrimination, and (c) strategic discrimination. In Section VI, we use these predictions to help determine which of these theories best explains the observed voting patterns. As discussed above, we focus this portion of our analysis on voting patterns in the Round 1. In addition, we employ tests that do *not* depend on whether contestants' optimal strategy is to vote off strong or weak players. Figure 1 summarizes the empirical tests discussed below and identifies those that rely solely on first round voting behavior.

Type of Discrimination	Empirical Prediction
1. Preference-Based Discrimination	Contestants will be more likely to vote against members of other groups.* Discrimination will become less pronounced as the game progresses since the implicit cost of discrimination increases at the end of the game.
 Statistical Discrimination a. Real group differences in average ability. 	There will be group-level differences in the percentage of questions answered correctly.* Contestants from all groups will be more likely to vote against members of a
b. Contestants have better information about the ability of members of their own group.	Contestants will be more likely to vote against members of their own group.*
 c. Inaccurate prior beliefs about a group's average ability. 	There will be a weaker pattern of discrimination in shows in which contestants are drawn from the same occupation.*
3. Strategic Discrimination a. Explicit collusion	Votes of members of a particular group will be correlated even after controlling for the votes of nongroup members. This also tests for implicit collusion based on
b. Implicit collusion	observable characteristics.* If a group successfully votes off a nongroup member in one round, then the group will be more likely to vote against a nongroup member in the subsequent round.
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Figure 1 Tests for Discrimination

A. Statistical Discrimination

Contestants in this game may use race and gender as proxies for underlying ability. In this setting, discriminatory voting patterns can arise either because of real group differences in ability or because some contestants have better information about the ability level of one group relative to another. These two sources of statistical discrimination (differences in ability vs. differences in the quality of information) yield different predictions about voting patterns.

To see this, assume that there are two groups: Group A and Group B, and that contestants from Group A play the game better than contestants from Group B. If the incentives are such that players wish to remove weak players, then contestants from *both* groups should be more likely to vote against members of Group B. Likewise, if the incentives are to vote off stronger players, then both groups should be more likely to vote against members of Group A. One simple method for examining whether this type of statistical discrimination accounts for observed voting patterns is to look for group-level differences in the percentage of questions answered correctly during the first round of the game. Further, if there are real group differences in average ability, then we would expect members of all groups to vote against members of a single group.

Now, suppose instead that contestants from Group A and Group B are equally skilled, but contestants are better able to evaluate the ability of contestants within their own group. For example, in the context of the show, women may be better able to assess whether or not another woman should be able to answer a given question correctly. In this case, contestants will place a relatively high weight on average ability when assessing the ability of players from other groups. That is, players from Group A will perceive a contestant from Group B with a below-average performance in Round 1 to be of *higher* ability than a contestant from Group A with that same performance. As a result, contestants will be *more* likely to cast votes *against* members of their own group. This occurs regardless of whether or not the incentives are to vote off strong or weak players are within their own group. Thus, information is harmful. The exact opposite occurs in hiring practices where workers from groups that emit noisy productivity signals are relatively unlikely to be hired. A final possibility is that the contestant's prior beliefs about ability are simply inac-

A final possibility is that the contestant's prior beliefs about ability are simply inaccurate. It is possible, for example, that members of Group A begin the game with the erroneous belief that members of Group B are inferior at playing the game. Thus, if contestants optimally vote against weak players, we would expect members of Group B to be more likely to be voted off than members of Group A even in the absence of average differences in group performance. Two types of analysis allow us to examine the possibility that players have such erroneous expectations. First, we examine data from episodes of the daily show for which all participants have the same occupation. Presumably members of the same occupation have better information about how particular groups will perform. If bad information is the reason that particular groups are voted off, one would expect to see a weaker pattern of discriminatory voting in shows where everyone has the same occupation. Second, if incorrect prior beliefs are driven by a characteristic that is not common to all members of the discriminated group, then the test for explicit collusion discussed below also can be used to rule out incorrect prior beliefs as a source of discriminatory voting behavior. A formal derivation of our predictions regarding statistical discrimination is presented in the Appendix.

B. Strategic Discrimination

In a strategic setting such as *The Weakest Link*, collusion also may play a role in discriminatory outcomes and may be the mechanism through which contestants either statistically discriminate or exhibit discriminatory preferences. Statistical discrimination can take this form when it is easier to form collusive agreements within one's own group, and preference-based discrimination can take this form when contestants prefer own-group agreements to cross-group agreements. Two types of collusion are possible here: explicit collusion (where particular contestants follow some agreed upon pattern of voting) and implicit collusion (where no agreements are discussed but individuals play strategies that yield focal points).

Contestants meet one another prior to the game and presumably have opportunities to discuss (quite possibly collusive) strategies for playing the game. These collusive agreements may be more likely to occur among members of the same race or gender. If contestants are explicitly colluding, they should not only be more likely to vote against nongroup members, but also should be more likely to vote against the *same* nongroup member. To test for explicit collusion, we examine whether the votes of the contestants are correlated with the votes of contestants from their own group. Because there may be unobservable characteristics that make particular individuals more likely to be voted off, in this analysis we also control for the votes cast by nongroup members.

If we find that votes are no more correlated within a group than outside the group, then we are also able to rule out many other forms of discrimination. Namely, if one group is discriminating based upon an unobservable individual level characteristic that is correlated with membership in another group, then this too will lead to correlation in within group voting.

Discriminatory outcomes also may result from implicit collusion where individuals naturally play particular equilibria. For example, discriminating against members of Group B may be a best response for members of Group A given that the other members of Group A are playing a discriminatory equilibrium. These focal point equilibria should be reinforcing. Hence, if Group A succeeds in removing a member of Group B in Round 1, then this should make it more likely that they will vote against a member of Group B in Round 2. We can therefore test for implicit collusion by examining how who is voted off in one round affects voting behavior in the next round. Note that testing for implicit collusion based upon an unobservable feature correlated with group membership is embedded in the test for explicit collusion. This is because, given this style of implicit collusion, within group votes will be correlated after controlling for the votes of others and the group membership.

C. Preference-Based Discrimination

In the prototypical model of preference-based discrimination, members of the majority group simply dislike working with members of the minority group and act as if there is some nonmonetary cost associated with hiring them. As a result, in equilibrium, workers from minority groups will earn lower wages than workers from preferred groups. In *The Weakest Link*, contestants from one group may dislike playing the game with contestants from some other group. Thus, if we believe that members of Group A dislike members of Group B, we would expect that members of Group A would be consistently more likely to vote off members of Group B in every round. As the game progresses, however, the probability that any single player will win the game increases. As a result, the implicit cost of preference-based discrimination *also* increases. For this reason alone, even if the contestants have discriminatory preferences, their propensity to discriminate will fade as the game progresses. Therefore, if preference-based discrimination exists, we would expect it to diminish with each round of the show.

V. Which Contestants Discriminate?

A. Descriptive Evidence

In order to understand the broad patterns in the data, we start by analyzing voting patterns by gender and race. Table 1 summarizes the voting behavior by round for both the daily and weekly shows. The table evaluates the voting behavior of three demographic groups—men, women, and whites—to determine if members of these groups discriminate against players who are not group members. For example, do men discriminate against women in their voting patterns? The voting behavior of blacks is not considered due to the small number of episodes with two or more black contestants.

The task of discerning discrimination is complicated because even if both men and women vote randomly, men will cast more votes against women and women will cast more votes against men simply because a contestant will never vote against him or her self. To account for this problem, and the fact that the distribution of demographic types varies across episodes and rounds, we describe the voting behavior in terms of the mean "group bias statistic." For individual *i* voting in round *r*, the "group bias statistic" is given by:

(1) GroupBiasStatistic =
$$\frac{l_{ir}}{\left[\frac{G_{ir}-1}{N_r-1}\right]}$$
.

where 1_{ir} is an indicator variable that takes on a value of one if individual *i* votes against a contestant from his or her group in Round *r*. G_{ir} is the number of contestant *i*'s type in Round *r* of his or her episode and N_r is the total number of contestants in Round *r* of contestant *i*'s episode. A mean value of one for the group bias statistic implies no discrimination, a value less than one implies discrimination against the other group and, and a value greater than one implies discrimination against one's own group.

The descriptive evidence shows some surprising results. First, there is virtually no evidence of discrimination by whites against blacks; all of the values of our discrimination statistic are indistinguishable from one. Second, there is no discrimination against women by men in the early rounds, though in Round 3 of both shows men are more likely to vote against women.⁹ The most surprising result, however, is that in

^{9.} This discriminatory behavior by men in round three loses statistical significance once we control for other factors and may be the result of retaliation or selection due to the discrimination by women against men in the early rounds of the game.

		Females			Males			Whites	
Show/Round	Sample	Mean	Standard Deviation	Sample	Mean	Standard Deviation	Sample	Mean	Standard Deviation
Daily 1		0.730*	1 1 1 1		9200	000	221	0.040	0 634
7	189	0.783*	1.141	186	1.336	0.336	235	0.982	0.703
l m	150	0.910	1.177	123	0.756^{*}	1.205	119	0.966	0.856
4	99	606.0	1.003	62	0.935	1.006	68	1.000	1.007
Weekly									
1	111	0.783*	1.086	113	1.000	1.185	133	1.004	0.479
2	98	0.907	1.119	98	0.978	1.344	103	1.019	0.526
n	84	0.933	1.160	83	0.689^{*}	1.172	82	1.037	0.500
4	64	1.115	1.307	72	0.806	1.162	62	0.871	0.689
S,	4	0.784	0.252	53	1.301	1.349	32	0.703	0.851
9	26	1.077	0.287	22	1.091	1.019	12	0.833	1.030
Note: The "group bias statistic" is equal to $l_{ir}/(N_r - 1)(G_{ir} - 1)$. Where l_{ir} is an indicator variable that takes on a value of one if individual <i>i</i> votes against a contestant from his/her group in round <i>r</i> . G_{ir} is the number of contestant \hat{r} 's type in round <i>r</i> of his/her episode and N_r is the total number of contestants in round <i>r</i> of contestant \hat{r} 's type in round \hat{r} of his/her episode and N_r is the total number of contestants in round <i>r</i> of contestant \hat{r} 's type in round \hat{r} of his/her episode and N_r is the total number of contestant is not contestant \hat{r} 's type in round \hat{r} of his/her episode and N_r is the total number of contestant in round \hat{r} of contestant \hat{r} 's type in round \hat{r} of his/her episode and N_r is the total number of contestant \hat{r} such that \hat{r} is the total number of contestant \hat{r} is th	is equal to $l_{ir}/($	$N_r - 1)(G_{lr} - 1)$ of contestant \vec{r} s	. Where 1_{ir} is an i type in round r of	indicator varial f his/her episoo	ble that takes of N_r is the and N_r is the second se	on a value of one e total number c	e if individual of contestants	<i>i</i> votes again in round <i>r</i> of	ist a contestant contestant i's
episode. A mean value of one implies no discriminatory outcomes, a value less than one implies discrimination against the other group and a value greater than one	nplies no discrim	inatory outcome	es, a value less tha	un one implies	discrimination	against the oth	er group and a	i value greate	r than one

Own-Group Bias Statistic

Table 1

implies discrimination against one's own group. The statistic is only calculated for rounds in which contestant i has the option of voting against both group members and nongroup members. * Statistically different from 1 at the 95 percent confidence level.

Round 1 of both the weekly and the daily show women are more likely to vote against men than women. This pattern of women voting off men continues in Rounds 2 and 3 of the daily show, though at a diminishing rate.

Obviously, these patterns may reflect race and gender differences in the average ability. In particular, if men are not as successful as women at playing the game, then this may explain why they are more likely than women to receive votes in the early stages of the game.

B. Evidence from Conditional Logits

In order to control for race and gender differences in the contestants' abilities to play the game, we estimate conditional logits by show and by round of the probability contestants cast votes against other players as a function of those players' characteristics. We model the utility of player i voting against contestant j in round r as:

(2)
$$U_{ijr} = X_{ijr}\beta_r + \varepsilon_{ijr}$$

where β is a vector of coefficients to be estimated and ε_{ijr} is the unobserved preference individual *i* has for voting against Contestant *j*. Included in X_{ijr} are controls for the percentage of questions the player answered correctly in that round, whether the player was the weakest link in that round and also whether the player was the strongest link in that round. In addition, we control for the gender and race of the other contestants and whether the individual voting was of the same gender and race. Finally, for all rounds other than Round 1, we include cumulative percent correct across all rounds and an indicator for whether the contestant doing the voting (Contestant *i*) ever received a vote from a contestant he is now considering voting against (Contestant *j*). This is to control for past performance and well as the possibility of retaliatory votes. Due to small sample sizes, we restrict our analysis of race to the daily show, though none of the qualitative results change if race is included in the weekly show. We assume that the ε_{ijr} 's are distributed i.i.d. extreme value, implying that the probability of voting against contestant *j* in round *r* is given by:

(3)
$$P_{ikr} = \frac{\exp(X_{ijr}\beta_r)}{\sum_{k=1}^{N_r} \exp(X_{ikr}\beta_r)}$$

where N_r is the number of contestants in round r.

The advantage of using a conditional logit is threefold. First, consider the probability that Player *i* votes against Player *j*. As the abilities of the other contestants increase, so too will the probability of voting against Player *j*. That is, the characteristics of the all the contestants influence the probability of voting against any one contestant. Second, the conditional logit allows us to examine interactions between the voting characteristics of the individual and the other contestants. Hence, we allow men to treat women differently *and* women to treat men differently. Third, the predicted number of votes cast by each contestant is constrained to be one and therefore the total number of predicted votes cast in a round equals the number of contestants in that round.

Results for Round 1 of the daily show and the weekly show are given in Table 2.

Results for the remaining rounds are shown in the Appendix. We have the most confidence in the Round 1 results for three reasons. First, sample sizes are larger in Round 1 than in the later rounds. Second, strategies may change as the game progresses and may depend upon the history of play. Finally, the pool of contestants in all but the initial round is endogenously determined by the voting behavior.

The first panel of Table 2 examines the voting behavior of women. Consistent with the descriptive evidence, women appear to be more likely to cast votes against men than against other women. This effect is quite large. Consider a daily show with six white contestants, three men, and three women, all with identical performance. The probability of a given woman voting against a particular man is 23.3 percent while the probability that she votes against a particular woman is only 15.1 percent. As shown in the Appendix, this effect disappears immediately after Round 1 in the weekly show while diminishing more slowly in the daily show. Also consistent with the descriptive evidence, the next two panels of Table 2 show that there is no indication of discrimination by men against women or by whites against blacks. This holds true for all rounds of the game.

As one would expect, for all demographic groups in the early stages of the game, the higher the percentage of the questions that the player answers correctly the less likely other contestants are to cast votes against that player. However, as the game progresses, the percent correct becomes less and less important. This confirms the basic logic that players have an incentive to vote against weak players in the early rounds of the game but, in the later rounds, this incentive is partially offset by the incentive to vote against stronger players.

VI. Why Women Vote Against Men

This section attempts to distinguish between the three possible hypotheses for why women are more likely to vote against men: statistical discrimination, strategic discrimination, and preference-based discrimination. Here again, we primarily focus our analysis on voting behavior in Round 1.

A. The Case Against Statistical Discrimination

Recall that statistical discrimination with correct priors can take two forms. In the first, the mean performance level is different across groups. Assuming that players wish to vote off weak players in Round 1, then if statistical discrimination of this type explains women's voting patterns, the performance of males must be on average worse than the performance of females. Table 3 documents the average percent correct for males and females by round and by show type. There is virtually no difference in performance levels for males and females in either show in the early rounds, while in the later rounds males actually perform better than their female counterparts.¹⁰ In addition, if men are truly worse than women at playing the game, then both

^{10.} That males perform better in later rounds may be a result of low-ability males being voted off sooner than low-ability females due to the discrimination.

`	5			
		Femal	Females Only	
Characteristics of other contestants	Daily Show	how	Week	Weekly Show
Male Percent correct Weakest link Strongest link Black Same race Observations	0.433* -2.667* 0.305 -0.016 -0.145 -0.189 222	(0.164) (0.470) (0.242) (0.242) (0.242) (0.292) (0.291)	0.538* -2.913* 0.429 0.275 111	(0.224) (0.506) (0.283) (0.510)
		Male	Males Only	
Characteristics of other contestants	Daily Show	Show	Weekl	Weekly Show
Female Percent correct Weakest link Strongest link Black Same race Observations	0.099 -2.018* 0.085 -0.596 0.199 -0.201	(0.150) (0.410) (0.234) (0.332) (0.308) (0.309)	-0.084 -2.754* 0.354 -0.728 113	(0.210) (0.488) (0.263) (0.746)

 Table 2

 Conditional Logit Estimate of Round 1 Voting Behavior

	Whi	Whites Only
Characteristics of other contestants	Daily Show	Weekly Show
Female Same gender Percent correct Weakest link Strongest link Black Observations	-0.170 -0.271* -2.508* 0.057 -0.360 0.000 349	(0.124) (0.123) (0.352) (0.192) (0.192) (0.144)

Note: Conditional logit estimates of the probability of an individual voting against **extra data a particular contestant. Standard errors are in parenthesis. * Statistically different from zero at the 95 percent confidence level.

		Females	ales		M	Males		В	Blacks		M	Whites
Show/ Round	Sample	Percent Correct	Standard Deviation									
Daily												
.	225	0.668	0.306	225	0.667	0.340	97	0.639	0.346	353	0.675	0.317
0	186	0.630	0.341	184	0.632	0.336	79	0.657	0.325	291	0.624	0.342
n	159	0.600	0.308	141	0.614	0.337	63	0.583	0.366	237	0.613	0.309
4		0.580	0.353	110	0.659	0.292	48	0.647	0.334	177	0.611	0.325
Weekly												
1	111	0.608	0.367	113	0.631	0.328	27	0.629	0.391	196	0.618	0.342
7	98	0.613	0.311	98	0.611	0.344	23	0.563	0.343	173	0.618	0.325
e	85	0.579	0.303	83	0.631	0.302	20	0.698	0.319	148	0.592	0.299
4	67	0.569	0.274	73	0.655	0.308	18	0.580	0.323	122	0.619	0.292
S	49	0.511	0.252	63	0.639^{*}	0.246	12	0.636	0.277	100	0.577	0.254
9	40	0.539	0.287	4	0.548	0.236	8	0.413	0.289	76	0.558	0.254

* Statistically different from female performance in the round at the 95 percent confidence level.

Table 3

women *and* men should be more likely to vote against men in the early rounds of the game. We find no evidence of this. An analogous argument holds if players wish to vote off strong players in the first round. Hence, there is no evidence that statistical discrimination based upon differences in mean group performance is driving females to vote against males.

The second type of statistical discrimination, where the signals on ability are more informative for one group than another, is ruled out by the fact that women are discriminating against men rather than against women. Recall that information-based statistical discrimination implies that within-group performance is more informative than out-of-group performance. Hence, both poor performance and good performance are weighted more heavily by members of one's own group—implying that women would be more likely to vote against women than against men. We find the exact opposite result.

Next, we turn to the case of incorrect prior beliefs regarding group performance. It is possible that in the early rounds of the game, women have not yet learned that men perform as well as women. Our data set contains 13 episodes of the daily show in which all of the contestants have the same occupation. Should erroneous prior beliefs exist, it seems reasonable to expect that workers in the same occupation would have more informative priors on the abilities of their opposite-sex contestants than do contestants of differing occupations. Thus, in Table 4, we reexamine women's voting behavior in Round 1 and interact a dummy variable for the shows in which all of the contestants have the same occupation with the male indicator. Under the incorrect priors hypothesis, one would expect a negative coefficient on this interaction term. However, the coefficient on this term is positive, although insignificant, providing weak evidence that women in the same occupation are more likely to vote against the hypothesis of incorrect prior beliefs in our discussion of explicit collusion below.

Table 4

Do Females Have Wrong Priors on Male Ability? Evidence from Same Occupation Shows

Characteristics of Other Contestants		ınd 1 es Only
Male	0.398*	(0.178)
Male \times same occupation show	0.238	(0.472)
Percent correct	-2.649*	(0.471)
Weakest link	0.309	(0.243)
Strongest link	-0.026	(0.342)
Black	-0.139	(0.292)
Same race	-0.192	(0.292)
Observations	222	

Note: Conditional logit estimates of the probability of a female voting against a particular contestant. Standard errors are in parentheses.

* Statistically different from zero at the 95 percent confidence level.

B. The Case Against Strategic Discrimination

We now test whether women are acting cooperatively with one another. There are two basic types of collusive behavior in which women might engage: explicit collusion and implicit collusion. Under explicit collusion, the presumption is that the women are following some agreed upon pattern of voting. Under implicit collusion, the presumption is that women are implicitly using men as focal points for collusive behavior.

To test for explicit collusion, we include in the female Round 1 conditional logits the total votes cast for each contestant by the other contestants as well as the total votes cast by the other female contestants.¹¹ The first variable captures the fact that there may be unobservable characteristics that lead *all* individuals to vote against a particular contestant. The second variable captures correlation among the votes of women—controlling for unobservable characteristics that draw the votes of all contestants. In order for there to be explicit collusion among females, the coefficient on this latter variable, total votes cast by other women, must be positive. This is true regardless of whether it is optimal for contestants to vote off strong or weak players. Results for this specification are shown in Table 5. Importantly, the total votes cast by other women has no more predictive power than the total votes cast by men. This can be seen because once we control for the total votes cast, the coefficient on votes cast by women is small and insignificant. Hence, there is no evidence of explicit collusion.

		Females Or	nly—Round 1	
Characteristics of Other Contestants	Daily	Show	Weekl	y Show
Male	0.470*	(0.200)	0.480*	(0.242)
Percent correct	-1.920*	(0.504)	-2.312*	(0.551)
Weakest link	0.156	(0.262)	0.243	(0.300)
Strongest link	0.028	(0.354)	0.198	(0.512)
Black	-0.161	(0.321)		
Same race	-0.237	(0.321)		
Total votes cast by				
Other contestants	0.541*	(0.102)	0.223	(0.120)
Other female contestants	-0.054	(0.179)	-0.028	(0.206)
Observations	222	. /	111	. ,

Table 5 Are Female Votes Correlated? Testing for Explicit Collusion

Note: Conditional logit estimates of the probability of a female voting against a particular contestant. Standard errors are in parenthesis.

* Statistically different from zero at the 95 percent confidence level.

^{11.} We also performed this analysis for the male subsample and for Round 2 of the daily show. Gender again had no effect on the voting behavior of men and the results for Round 2 were very similar to the results from Round 2; female votes are not correlated except through total votes.

The fact that the coefficient on female votes is small and insignificant also allows us to reject many other explanations for the discriminatory behavior of women. In particular, it suggests that women are not discriminating on some unobserved characteristic correlated with being male as this too should have led to a positive coefficient on votes cast by women. For example, if particular men are arrogant or make poor (in the eyes of women) banking decisions, then the votes of the other women should reflect this. Instead, the results suggest that any unobservable characteristics of men that are unappealing to women are equally unappealing to men. Note that this also helps to rule out discrimination based upon bad expectations. In particular, the only bad expectations that can exist involve women expecting all men to perform equally poorly—that is, there can be no correlation between these bad expectations and any observable (to the voter, not to the researcher) characteristics.

We next test whether or not women are implicitly colluding to vote off men. The first evidence that this is not the case comes from the diminishing coefficient on same sex as we move to the later rounds of the game. If implicit collusion is working, then there is no reason to stop colluding against men in the later rounds. The second indication that implicit collusion is not driving the results is that implicit collusion does not appear to be reinforcing. That is, if implicit collusion successfully occurs in one round (a man is voted off), then it should be more likely to occur in the next round. Again, this prediction does not depend on whether or not the optimal strategy is to vote off strong or weak players. The top panel of Table 6 reports the conditional logits for Round 2 of the daily and weekly shows with an indicator variable for whether a man was voted off in Round 1. Similarly, the bottom panel reports the conditional logits for Round 3 with an indicator variable for whether a man was voted off in both Round 1 and Round 2. Although not significant, three of the four interactions are negative, implying that, if anything, removing males makes it less likely that women will vote off men in future rounds. Hence, we find no evidence of strategic discrimination through either implicit or explicit collusion.

These results should be interpreted with some caution. Given that there is selection into later rounds, contestants update their beliefs regarding the abilities of their opponents and observe the votes of their opponents in previous rounds. To partially control for these factors, we include the cumulative percent correct from the previous rounds as well as whether the opponent had ever voted against the contestant in any of the previous rounds. The discrimination coefficients for women—and correspondingly the discrimination coefficients for men and for whites—are not sensitive to the inclusion of these variables.

C. The Case for Preference-Based Discrimination

The only remaining explanation is that women simply prefer playing with women. Consistent with this explanation, the coefficient on male diminishes in later rounds as the price of discriminating based upon preferences increases. Further, this coefficient falls faster in the weekly show, disappearing after one round. This is consistent with the theory of preference-based discrimination since the cost of discriminating is higher in the weekly show where the total prize money is substantially larger.

We have attempted to further test this explanation by including controls for the amount of money banked (the size of the pot) at the voting stage. If preference-based

	1	Females On	ly—Round 2	2
Characteristics of Other Contestants	Daily S	Show	Weekly	y Show
Male	0.480	(0.338)	0.313	(0.418)
Male \times male voted off in Round 1	-0.285	(0.396)	-0.592	(0.518)
Percent correct	-1.526*	(0.455)	-3.551*	(0.673)
Past percent correct	-0.349	(0.327)	-0.833*	(0.422)
Weakest link	0.255	(0.255)	0.035	(0.323)
Strongest link	-0.071	(0.293)	-0.329	(0.786)
Voted against	1.405*	(0.338)	1.742*	(0.421)
Black	-0.070	(0.307)		· /
Same race	-0.129	(0.309)		
Observations	179	. ,	98	
]	Females On	ly—Round 3	}
Characteristics of other contestants	Daily S	Show	Weekly	y Show
Male	0.160	(0.251)	0.225	(0.299)
Male \times male voted off in	0.030	(0.412)	0.654	(0.637)
Rounds 1 & 2				· /
Percent correct	-1.584*	(0.518)	-2.942*	(0.714)
Past percent correct	-1.171*	(0.536)	-1.639*	(0.694)
Weakest link	0.061	(0.138)	0.323	(0.337)
Strongest link	-0.023	(0.308)	-0.156	(0.600)
Voted against	1.232*	(0.314)	0.904*	(0.438)
Black	-0.722	(0.559)		
Same race	-1.070	(0.556)		
Observations	156		85	

Table 6

Do Females Use Males as Focal Points? Testing for Implicit Collusion

Note: Conditional logit estimates of the probability of a female voting against a particular contestant. Past percent correct is the cumulative percent correct in the previous rounds. Voted against is whether the opponent ever voted against the contestant in any of the previous rounds. Excluding past percent correct and voted against does not change the coefficients on the male variables. Standard errors are in parenthesis. * Statistically different from zero at the 95 percent confidence level.

discrimination exists, higher amounts of money banked should lead to less discrimination. Estimates of models of this type were mixed, with evidence that at the lowest quartile of money banked discrimination increases (consistent with the theory) but discrimination also increases at the highest quartile of money banked (inconsistent with the theory). One possible explanation for the latter is that, in order to bank a large sum of money, every contestant needs to perform well, and if every player is similarly talented at playing the game, then there is no extra cost associated with voting against men, regardless of the amount of money banked. Unfortunately, the small sample sizes make it difficult to further test this hypothesis.

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Characteristics of Other Contestants	Rot	Round 1	Round 2	1 2	Round 3	nd 3	Round 4	d 4
Females only								
Male	0.433*	(0.164)	0.276	(0.176)	0.170	(0.202)	0.133	(0.291)
Percent correct	-2.667*	(0.470)	-1.538*	(0.453)	-1.584^{*}	(0.518)	-0.775	(0.809)
Past percent correct		r.	-0.346	(0.327)	-1.175	(0.533)	-0.453	(0.913)
Weakest link	0.305	(0.242)	0.246	(0.253)	0.060	(0.138)	0.195	(0.380)
Strongest link	-0.016	(0.342)	-0.070	(0.293)	-0.022	(0.307)	0.405	(0.340)
Voted against			1.402*	(0.338)	1.232*	(0.314)	1.232*	(0.439)
Black	-0.145	(0.292)	-0.065	(0.307)	-0.718	(0.557)	-0.798	(0.585)
Same race	-0.189	(0.291)	-0.127	(0.308)	-1.067	(0.555)	-1.322*	(0.596)
Observations	222		179		156		113	
Males only								
Female	0.099	(0.150)	-0.09	(0.171)	0.278	(0.236)	-0.056	(0.281)
Percent correct	-2.018*	(0.410)	-1.843*	(0.445)	-2.289*	(0.625)	-0.786	(0.809)
Past percent correct			-0.493	(0.343)	-0.761	(0.550)	-0.486	(0.944)
Weakest link	0.085	(0.234)	-0.033	(0.250)	0.024	(0.173)	0.188	(0.366)
Strongest link	-0.596	(0.352)	-0.395	(0.320)	-0.191	(0.360)	0.232	(0.367)
								(Continued)

 Table 7

 Conditional Logits of Voting Behavior in the Daily Show

Characteristics of Other Contestants	Rot	Round 1	Round 2	nd 2	Rou	Round 3	Round 4	d 4
Voted against			1.369*	(0.307)	1.106*	(0.341)	0.899*	(0.419)
Black	-0.199	(0.308)	-0.026	(0.314)	-0.347	(0.453)	-0.504	(0.484)
Same race	-0.201	(0.309)	0.033	(0.317)	0.188	(0.453)	-0.303	(0.489)
Observations	222		179		138		110	
Whites only								
Female	-0.170	(0.124)	-0.161	(0.138)	0.007	(0.167)	-0.290	(0.228)
Same gender	-0.271^{*}	(0.123)	-0.131	(0.138)	-0.241	(0.165)	0.118	(0.217)
Percent correct	-2.508^{*}	(0.352)	-1.918^{*}	(0.359)	-1.765*	(0.473)	-0.605	(0.624)
Past percent correct			-0.589^{*}	(0.266)	-0.863*	(0.415)	-0.320	(0.734)
Weakest link	0.057	(0.192)	-0.076	(0.202)	-0.025	(0.161)	0.245	(0.304)
Strongest link	-0.360	(0.284)	0.006	(0.234)	-0.214	(0.255)	0.295	(0.270)
Black	0.000	(0.144)	-0.012	(0.165)	-0.053	(0.200)	0.107	(0.266)
Voted against			1.433^{*}	(0.253)	1.192^{*}	(0.255)	1.170^{*}	(0.336)

rounds. Voted against is whether the opponent ever voted against the contestant in any of the previous rounds. Standard errors are in parenthesis.

* Statistically different from zero at the 95 percent confidence level

 Table 7 (continued)

Characteristics of Other Contestants	Round	ind 1	Rou	Round 2	Rou	Round 3	Round 4	ıd 4	Round 5	ıd 5	Rou	Round 6
Females only												i t
Male Derrent correct	0.538*	(0.224)	-0.060 -3.612*	(0.246)	0.371	(0.265)	00.0 880.0	(0.302)	0.050	(0.408)	-0.994	(0.746)
Past percent		(000-0)	.710.6-	(7/0.0)		(017.0)	0.000	(001.0)	1/0.0	(160.1)	C17.1	(060.1)
Correct			-0.853*	(0.420)	1.503*	(0.673)	2.329*	(1.050)	-1.050	(1.381)	5.499	(3.286)
Weakest link		(0.283)	-0.007	(0.317)	0.328	(0.339)	0.952*	(0.401)	0.256	(0.469)	2.204*	(0.931)
Strongest link	0.275	(0.510)	-0.325	(0.787)	-0.155	(0.597)	-1.045	(0.535)	-0.500	(0.487)	0.227	(0.653)
Voted against		1.774*	(0.420)	0.922^{*}	(0.436)	1.168^{*}	(0.387)	1.633^{*}	(0.606)	.035*	(1.163)	
Observations	111		98		85		99		49		39	
Males only		(010.0)	0155	(0100)	0 457			(100.0)	0 220		0400	() E04)
remale	-0.084	(017.0)	CC1.0	(0.243)	0.455	(607.0)	-11.0-	(175.0)	PCC.0-	(075.0)	-0.488	(400.0)
Percent correct	-2.754*	(0.488)	-2.818^{*}	(0.606)	-2.560*	(0.700)	-2.476*	(0.855)	-0.087	(1.177)	1.535	(1.549)
Correct			-1.538	(0.448)	-2.167*	(0.717)	$\frac{2.845}{}$	(1.092)	0.190	(1.245)	1.627	(2.521)
Weakest link	0.354	(0.263)	-0.027	(0.330)	0.383	(0.319)	0.311	(0.343)	0.357	(0.426)	1.300*	(0.625)
Strongest link	-0.728	(0.746)	-34.200		-0.293	(0.655)	-0.712	(0.611)	-0.227	(0.444)	0.286	(0.579)
Voted against			1.192	(0.431)	0.359	(0.433)	1.292*	(0.460)	1.136^{*}	(0.386)	0.115	(0.567)
Observations	113		98		83		71		63		42	

 Table 8

 Conditional Logits of Voting Behavior in the Weekly Shows

Note: Conditional logit estimates of the probability of an individual voting for a particular contestant. Past percent correct is the cumulative percent correct in the previous rounds. Voted against is whether the opponent ever voted against the contestant in any of the previous rounds. No males voted for the strongest link in Round 2 of the weekly *Statistically different from zero at the 95 percent confidence level. show. Standard errors are in parenthesis.

VII. Discussion

Given that we are interested in learning about labor market discrimination, it is important to confirm that our findings are consistent with what we observe in other contexts. A number of recent audit studies (for example, Goldin and Rouse 2000 and Bertrand and Mullanaithan 2004) have documented evidence of discrimination against women and blacks. We find no evidence of discrimination against these groups.

There are a number of explanations for this discrepancy. First, the contestants on the *The Weakest Link* operate under the scrutiny of a much larger audience than employers in the labor market. Thus, if there is a social stigma associated with discrimination (of any kind), then individuals may not be willing to discriminate when their actions are publicly observable. This suggests that open hiring policies may reduce discrimination. To our knowledge there has been no previous research on the impact of audiences on discriminatory behavior.

Second, as Table 3 shows, there are no significant performance differences between men and women or between blacks and whites in either the first or the second moment of the ability distribution. As a result, even if statistical discrimination is an important feature of the labor market, it would be unlikely to appear in our analysis. Thus, one interpretation of the fact that we find no evidence of discrimination against either blacks or women in our data is as support for the role of statistical discrimination in the discriminatory patterns found in recent studies of the labor market.

Somewhat surprisingly, we also find evidence that women have discriminatory preferences against men. There are a number of possible explanations for this type of behavior. First, women may dislike certain aspects of how men play the game.¹² Women may also feel more compassionate toward women than toward men, and women may not like playing with men because they fear that they will not compete as well against men in the later rounds of the game. In experimental settings, Gneezy, Niederle and Rustichini (2003) and Gneezy and Rustichini (2004), for example, find evidence that competition improves the performance of men but does not do so for women.

Evidence from other settings supports the notion that women might give preferential treatment to other women. As discussed previously, Dillingham, Ferber, Hamermesh (1994) find that women discriminate in favor of women in voting for officers in a professional association. There is also evidence that, all else equal, women are more likely to vote for women in political contests. For example, in her analysis of voting behavior during the 1992 U.S. election, Dolan (1998) finds that women were more likely than men to vote for women candidates, even after controlling for a number of ideological, issue, and party concerns. Similar results are reported by Smith and Fox (2001) for open seat house races between 1988 and 1992. Further, these authors find candidate sex does not matter to male voters once controls for other factors are

^{12.} Though for this to be true, it would have to be how males in general play the game and not correlated with any other feature (such as how certain males speak or how certain males look). If the discrimination result was driven by females disliking the way males with particular features played the game, the test we performed for explicit collusion would have shown that female votes were correlated beyond the correlation with male votes.

included. Finally, Derose et al. (2001) examine the relationship between patient satisfaction and the gender of emergency department physicians. They find that even after controlling for the patient's age, health status, literacy level, and a number of other covariates, women gave significantly higher performance ratings to female physicians. Men's satisfaction, on the other hand, was not associated with physician gender. While beyond the scope of this paper, understanding more about the sources of women's preferences toward women is clearly an intriguing area for future research.

VIII. Conclusion

Understanding the nature of discrimination and its contribution to both racial and gender earnings inequality involves tackling two questions. First, we would like to know whether individuals discriminate. Second, we would like to know why individuals discriminate. In this paper we attempt to address both of these questions by examining the voting behavior of contestants in *The Weakest Link*. Although the game show environment is clearly different from that of the labor market, difficulties associated with evaluating discrimination directly within the labor market motivate an analysis of behavior in more stylized settings. This research builds on the emerging experimental literature on discrimination and provides an opportunity to consider individual behavior in a high-stakes environment.

Interestingly, we find no evidence of either preference-based or strategic discrimination against either blacks or women. Statistical discrimination was ruled out in *The Weakest Link* in part because of the lack of performance differences across racial and gender lines. However, this may not be the case in the labor market and points toward statistical discrimination as a possible explanation for the discriminatory patterns against these groups found in a number of recent studies. In addition, we find that women discriminate against men in the early rounds of both the daily and weekly shows. The one theory consistent with the observed voting trends is preference-based discrimination. In other words, it appears that women simply prefer playing with women rather than with men. Finally, our paper suggests the need for future studies that examine the impact of audiences on discriminatory behavior.

Appendix

Here we present a simple model of statistical discrimination and discuss its implications for voting behavior in the first round of *The Weakest Link*. An essential feature of the model is that players base their voting decisions on an indicator of player ability, x, that approximates true ability, θ .

We allow contestants to belong to one of two groups: group *A* and group *B*. Since these groups may vary in their average ability, we let the distribution of θ depend on group membership. In particular, for group *A*, $\theta \sim N(\mu_A, \sigma_{\theta A}^2)$, and, for group *B*, $\theta \sim$ $N(\mu_B, \sigma_{BA}^2)$. Contestants, however, only observe $x = \theta + \varepsilon$, where ε represents a random disturbance. For example, *x* may represent an index of a player's overall performance in Round 1 including the number of questions that he or she answered correctly, the time it took him to answer those questions and the difficulty of the questions that he or she was asked. We assume that ε is distributed $N(0, \sigma_{\varepsilon}^2)$. Below, we develop the implications of this simple model. In particular, we first examine voting behavior when players correctly perceive that there are real group differences in average ability. Second, we examine voting behavior when players are better able to assess the ability signals of contestants from their own group. Finally, we consider what happens if players have incorrect prior beliefs about the distribution of θ . Interestingly, the model's empirical predictions do not depend on whether contestants wish to vote off strong players or weak players.

A. Baseline Model

In this model, contestants make inferences about θ given *x* using Bayes's Rule. Under our assumptions, the posterior distribution of θ for a contestant from group *j* with first round performance *x* is known to be normal with mean *m_i* and variance *s_i*, where

$$m_j = s_j x + (1 - s_j) \mu_j$$

and

$$s_j = \frac{\sigma_j^2}{\sigma_j^2 + \sigma_\epsilon^2}$$

Note that m_j is a weighted average of the player's Round 1 performance, x, and the prior mean, μ_j , where the weights depend on how diffuse is the prior on θ and how informative is the signal x.

B. Real Group Differences in Average Ability

Suppose that contestants from Group *A* have higher average ability than contestants from Group *B*, but the prior variance of θ is the same for both groups. That is, suppose that $\mu_A > \mu_B$ and $\sigma_A^2 = \sigma_B^2$. In this case it is clear that, given identical signals (realizations of *x*), m_A will be greater than m_B . Thus, if contestants wish to vote off weak players, then contestants will be more likely to vote off members of Group *A* than Group *B* and if contestants wish to vote off strong players, contestants will be more likely to vote off members of Group *A*. In either case, the central prediction is that contestants from *all* groups will be more likely to vote against members of a single group. In addition, to the extent that we are able to observe *x*, we also should observe average group differences in performance.

C. Asymmetric Information

Now suppose instead that contestants from Group A and Group B have the same average ability, but contestants are better at assessing the ability of contestants from their own group than they are of those from another group. For example, this would be the case if women (men) are better able to identify the types of questions that women (men) should be able to answer correctly. In the context of the model, we assume that $\mu_A = \mu_B$ and $\sigma_A^2 = \sigma_B^2$. Differences in the ability to assess ability are captured through the variance of ε . In particular asymmetric information is represented by the case where σ_{ε}^2 is lower when receiving a signal for a member of your own group and higher when receiving a signal from a member of another group. In this case, it is clear that s_j is higher for members of a participant's own group than for members of another group—implying that the posterior mean of θ will depend more heavily on xfor contestants from a players own group than it will for contestants from another group. Thus, players with both the highest and the lowest posterior mean ability will be from the contestant's own group. As a result, contestants will be more likely to vote against members of their own group than against contestants from another group, regardless of whether it is optimal for contestants to vote off strong or weak players.

D. Incorrect Prior Beliefs

Typically in models of statistical discrimination it is assumed that people have correct prior beliefs about the distribution of true ability. However, it may be possible that contestants on this program have inaccurate prior beliefs. For example, even if the distribution of θ is identical for men and women, women may incorrectly perceive that men are worse at playing the game than women. Since these beliefs are incorrect, they are naturally hard to verify. However, if incorrect priors explain discrimination, we would expect to see more accurate prior beliefs—and hence less discrimination—in games in which all contestants are drawn from the same occupation.

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